# No eigenvalues outside the support of limiting empirical spectral distribution of a separable covariance matrix

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#### Abstract

We consider a class of matrices of the form  $C_n=(1/N)A_n^{1/2}X_nB_nX_n^*A_n^{1/2}$ , where  $X_n$  is an  $n\times N$  matrix consisting of i.i.d. standardized complex entries,  $A_n^{1/2}$  is a nonnegative definite square-root of the nonnegative definite Hermitian matrix  $A_n$ , and  $B_n$  is diagonal with nonnegative diagonal entries. Under the assumption that the distribution of the eigenvalues of  $A_n$  and  $B_n$  converge to proper probability distributions, as  $\frac{n}{N}\to c\in(0,\infty)$ , the empirical spectral distribution of  $C_n$  converges a.s. to a non-random limit. We show that, under appropriate conditions on the eigenvalues of  $A_n$  and  $B_n$ , with probability 1, there will be no eigenvalues in any closed interval outside the support of the limiting distribution, for sufficiently large n. The problem is motivated by applications in spatio-temporal statistics and wireless communications.

**Keywords**: Empirical spectral distribution, Stieltjes transform, separable covariance, CDMA, MIMO.

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### 1 Introduction

The aim of this paper is to extend the result in Bai and Silverstein (1998) to the eigenvalues of a more general class of random matrices, specifically matrices of the form

$$C_n = (1/N)A_n^{1/2}X_nB_nX_n^*A_n^{1/2},$$

where for  $n=1,2,\ldots,X_n$  is  $n\times N$  (N=N(n)) consisting of i.i.d. standardized complex entries  $(\mathbb{E}X_{11}=0,\mathbb{E}|X_{11}|^2=1),\ A_n^{1/2}$  is a nonnegative definite square root of the  $n\times n$  Hermitian nonnegative definite matrix  $A_n$ , and  $B_n=\mathrm{diag}(b_1,b_2,\ldots,b_N)$ , each  $b_i\geq 0$ . The matrices studied in Bai and Silverstein (1998) assume  $B_n=I_N$ , the  $N\times N$  identity matrix. In that case  $C_n$  can be viewed as the sample covariance matrix consisting of N samples of the random vector  $A_n^{1/2}X_{-1}$   $(X_{-1}$  denoting the first column of  $X_n$ ), which has population covariance matrix  $A_n$ . The matrix  $C_n$  can then be interpreted as the sample covariance matrix consisting of N weighted samples. There are other ways to interpret the matrix, being important in various applications. One example is the spatio-temporal sampling model to be described in Section 1.2.1. In wireless communications,  $H_n=(1/\sqrt{N})A_n^{1/2}X_nB_n^{1/2}$ , for general nonnegative definite matrix  $B_n$ , is used to model the path gains between different groups of antennas in a multiple-input-multiple-output (MIMO) system (Section 1.2.2). It is typically assumed that  $X_{11}$  is complex Gaussian (real and imaginary parts independently distributed as N(0,1/2)), in which case the square of the singular values of  $H_n$  has the same distribution as the eigenvalues of  $C_n$  (the  $b_i$ 's being the eigenvalues of  $B_n$ ).

#### 1.1 Statement of the result

Results have previously been obtained on the limiting behavior of the empirical distribution function,  $F^{C_n}$ , of its eigenvalues ( $F^{C_n}(x) \equiv \text{(number of eigenvalues of } C_n \leq x)/n$ ), (Burda (2005), Zhang (2006), Boutet de Mondvel, Khorunzhy and Vasilchuck (1996)), with differing assumptions (the weakest appearing in Zhang (2006)) and varied (but equivalent) forms of expressions for the result. The following limit result is expressed in terms of the Stieltjes transform of  $F^{C_n}$ , defined for any distribution function G as

$$m_G(z) = \int \frac{1}{\lambda - z} dG(\lambda), \quad z \in \mathbb{C}^+ \equiv \{z \in \mathbb{C} : \Im z > 0\}.$$

Assume that the empirical distribution functions,  $F^{A_n}$  and  $F^{B_n}$ , converge weakly, as  $n \to \infty$ , to probability distribution functions, denoted respectively by  $F^A$  and  $F^B$ , and  $c_n \equiv n/N \to c > 0$ . Then, with probability 1,  $F^{C_n}$  converges weakly to a probability distribution function F whose Stieltjes transform  $m(z) = m_F(z)$ , for  $z \in \mathbb{C}^+$ , is given by

$$m(z) = \int \frac{1}{a \int \frac{b}{1+cbe} dF^B(b) - z} dF^A(a),$$
 (1)

where e = e(z) is the unique solution in  $\mathbb{C}^+$  of the equation

$$e = \int \frac{a}{a \int \frac{b}{1 + cbe} dF^B(b) - z} dF^A(a). \tag{2}$$

It is remarked here that the result in Zhang (2006) covers arbitrary Hermitian nonnegative definite  $B_n$ . Moreover, the assumption of identical distribution of the entries of  $X_n$  is weakened to a Lyapunov-type condition.

As in Bai and Silverstein (1998), the purpose of this paper is to prove, with additional assumptions, the almost sure non-appearance of eigenvalues of  $C_n$  in any interval away from the origin and outside the support of F as  $n \to \infty$ . Before the result can be formally stated we need one more definition. Let  $F^{c_n,A_n,B_n}$  denote the distribution function whose Stieltjes transform is given by (1) replacing c,  $F^A$  and  $F^B$  with  $c_n$ ,  $F^{A_n}$  and  $F^{B_n}$ , respectively.

The following will be proven:

THEOREM 1: Assume the following:

- (a)  $X_{ij}$ , i, j = 1, 2, ..., are i.i.d. complex-valued random variables with  $\mathbb{E}X_{11} = 0$ ,  $\mathbb{E}|X_{11}|^2 = 1$ , and  $\mathbb{E}|X_{11}|^4 < \infty$ .
- (b) N = N(n) with  $c_n = n/N \to c > 0$  as  $n \to \infty$ .
- (c) For each n,  $A_n$  is  $n \times n$  Hermitian nonnegative definite, and  $B_n = diag(b_1, \ldots, b_N)$  is  $N \times N$ , each  $b_i \geq 0$ , satisfying  $F^{A_n} \xrightarrow{\mathcal{D}} F^A$ ,  $F^{B_n} \xrightarrow{\mathcal{D}} F^B$ , both limits being probability distribution functions.
- (d)  $||A_n||$  and  $||B_n||$ , the respective spectral norms of  $A_n$   $B_n$ , are bounded in n.
- (e)  $C_n = (1/N)A_n^{1/2}X_nB_nX_n^*A_n^{1/2}$ , where  $A_n^{1/2}$  any Hermitian square root of  $A_n$ ,  $X_n = (X_{ij})$ , i = 1, 2, ..., n, j = 1, 2, ..., N.
- (f) The interval [a,b] with a > 0 lies in an open interval outside the support of  $F^{c_n,A_n,B_n}$  for all large n.

Then,

$$\mathbb{P}(no\ eigenvalue\ of\ C_n\ appears\ in\ [a,b]\ for\ all\ large\ n)=1.$$

The applicability of Theorem 1 depends on finding a way to determine the intervals outside the support of  $F^{c_n,A_n,B_n}$ , as it exists for sample covariance matrices (Silverstein and Choi (1995)). In the latter case, the limiting Stieltjes transform m(z) has an explicit inverse z = z(m). It is straightforward to verify that a Stieltjes transform is increasing on intervals on the real line outside the support of its distribution function. Its inverse therefore exists on these intervals and is also increasing. Therefore plotting z(m) for m real, and locating on the vertical axis places where the inverse is increasing, yield intervals outside the support. There does not appear to be an explicit inverse for (1). Nevertheless, preliminary work indicates a way to determine an inverse of m(z) associated with an interval outside the support of the limiting spectral distribution. This has been established in the case of another class of random matrices (Dozier and Silverstein (2007)). Work in this area is currently being pursued.

#### 1.2 Motivation

Our results give information on the behavior of individual eigenvalues. Results describing only the limiting behavior of the empirical spectral distribution provide information on the proportion of eigenvalues falling in any interval. But these results do not rule out the possibility of o(n) eigenvalues scattered outside the support of the limiting empirical spectral distribution. The goal of our research is to establish that such a phenomenon does not occur for large enough n. Further research in our framework would allow for precise description of the location of the eigenvalues. In

particular, we expect that the results proved here will be key to proving certain phase transition phenomena observed in the context of sample covariance matrices with  $B_n = I_N$  and  $A_n$  having a few large isolated eigenvalues (Baik and Silverstein (2006), Baik, Ben Arous and Péché (2006), El Karoui (2006), Paul (2007)).

#### 1.2.1 Application to spatio-temporal statistics

The data model we are considering here arises in the field of spatio-temporal statistics, where the rows of the  $n \times N$  matrix  $U_n = A_n^{1/2} X_n B_n^{1/2}$  correspond to indices of spatial locations and the column indices correspond to points in time. This class of models is also known as the *separable covariance model*. This is because, under the assumptions made here on the entries of  $X_n$  (i.i.d., mean 0, finite fourth moment), the joint (space-time) covariance of  $U_n$ , viewed as an  $Nn \times 1$  vector consisting of the columns of the matrix  $U_n$  stacked on top of one another, is given by  $\Sigma_{U_n} = A_n \otimes B_n$ , where  $\otimes$  denotes the Kronecker product between matrices. Note that, if we further assume Gaussianity for the entries of  $X_n$ , then the joint distribution of  $U_n$  is  $N_{Nn}(0, A_n \otimes B_n)$ . Also, in that setting, we do not require  $A_n$  and  $B_n$  to be diagonal, but only that they are nonnegative definite. The interpretation of this covariance structure is that the entries of  $U_n$  are correlated in time (column), but the pattern of temporal correlation does not vary with location (row). In other words, there is no space-time interaction in the process.

One advantage of this model from a statistical estimation point of view is that, when N is large and n is comparatively small, so that  $\frac{n}{N} \to 0$  as  $n \to \infty$ , it is possible to get quite reliable estimate of  $A_n$  from the sample covariance matrix  $C_n = \frac{1}{N}U_nU_n^*$ . Indeed, in that setting, if moreover  $||A_n||$  is bounded above, it is not hard to verify that  $||C_n - \frac{1}{N}(\operatorname{tr} B_n)A_n|| \to 0$  a.s., as  $n \to \infty$ . So, the spectral properties of  $A_n$  can be recovered from that of the spectrum of  $C_n$ . Of course, the key questions we are addressing here relate to the situation where  $\frac{n}{N} \to c \in (0, \infty)$ . The behavior of the empirical spectrum in that setting is hitherto unknown.

The results and techniques presented in this paper may prove useful in this problem for a number of different reasons. A statistical problem related to such spatio-temporal processes is to understand the temporal variability of the spatial field. One of the approaches for understanding the temporal variability is to perform an eigen-analysis (in space) of the sample covariance matrix  $C_n$ . This is because, the weights of the different eigenvectors of  $C_n$ , in representing the columns of  $U_n$  (principal components scores), vary in time. These weights therefore capture the temporal variability of the orthogonal components (eigenvectors of  $C_n$ ) of the spatial process. The eigenfunctions thus obtained are usually referred to as empirical orthogonal functions (particularly in climatology, see, e.g. von Storch and Zwiers (1999)). Understanding the asymptotic behavior of the sample eigenvalues and eigenfunctions therefore is a relevant question, since, under the separable space-time model they give a set of orthogonal components, and their relative strengths, of the spatial variation of the process.

#### 1.2.2 Application to wireless communication

In wireless communications,  $H_n = (1/\sqrt{N})A_n^{1/2}X_nB_n^{1/2}$ , for a general nonnegative definite matrix  $B_n$ , appears in a variety of models, including both direct-sequence and multiple-carrier code-division multiple-access systems (Tulino and Verdú (2004), sections 3.1-3.2), and in multiple-input-multiple-output (MIMO) systems (Tulino and Verdú (2004), section 3.3). The importance of acquiring more detailed information on the singular values of  $H_n$  beyond what the limiting empirical distribution ((1), (2)) reveals, which has been primarily used to estimate capacity, is becoming more apparent.

For example, in Verdú (2002) an estimate of capacity requires knowledge of the largest singular value of  $H_n$  which Theorem 1 provides. Indeed, the a.s. convergence of the largest singular value of  $H_n$  is an analog of the corollary to Theorem 1.1 in Bai and Silverstein (1998) and readily follows from Theorem 1 using similar arguments. Another example is in MIMO systems, where  $H_n$  models the path gains between different groups of antennas. It is typically assumed that  $X_{11}$  is complex Gaussian (real and imaginary parts independent N(0,1/2)), in which case the square of the singular values of  $H_n$  has the same distribution as the eigenvalues of  $C_n$  (the  $b_i$ 's being the eigenvalues of  $B_n$ ). The matrices  $A_n$  and  $B_n$  are the covariances between the receiver and the transmitter antennas, respectively. They reflect the scenario involving these two groups of antennas, for example, their locations, and the nature of the interference encountered due to their surroundings. The singular values of  $H_n$ , or equivalently the eigenvalues of  $C_n$ , indicate several important properties of the communication scheme, due to the fact that any information on  $H_n$  yields ways to allocate the transmitted signal in an optimal way. For example, if there is a significant number of small eigenvalues, transmission can be achieved after performing a unitary transformation, on the left and/or the right side of  $H_n$ , resulting in a reduced number of virtual parallel antennas with little correlation between them. When the number of antennas is sizeable, knowledge of the eigenvalues of  $C_n$ , depending only on  $A_n$  and  $B_n$ , is gained to some extent from the limiting F. It yields the proper proportion of eigenvalues within any interval. However, Theorem 1 is a step toward knowing the location of all the singular values, which provides much more information. For example, it can ensure that no lone eigenvalue above or below the limiting support exists. The importance of Theorem 1 lies in the determination of spectral behavior of  $C_n$ entirely through  $A_n$  and  $B_n$ .

The essential portion of the proof of Theorem 1 will proceed in the following sections. The main tools used in the proof are properties of Stieltjes transform and bounds on the moments of martingale difference sequences. The results to be obtained here are analogous to those in sections 3-5 of Bai and Silverstein (1998), namely, we will show

$$\sup_{x \in [a,b]} |m_n(z) - m_n^0(z)| = o(1/(nv_n)) \quad \text{a.s.},$$
(3)

where

$$m_n = m_n(z) = m_{FC_n}(z) = (1/n)\operatorname{tr}(C_n - zI)^{-1}$$
 (4)

is the Stieltjes transform of the empirical distribution function of the eigenvalues of  $C_n$ ,

$$m_n^0 = m_n^0(z) = m_{F^{c_n, A_n, B_n}}(z)$$
 (5)

and  $z = x + iv_n$ , where  $v_n = \kappa n^{-1/140}$ ,  $\kappa$  an arbitrary positive constant (fixed for all n).

The steps needed to conclude Theorem 1 from (3) are identical to those in section 6 of Bai and Silverstein (1998), except for the fact that in the latter paper  $v_n = N^{-1/68}$ . In particular, following the same arguments one can prove that, if a', b' are such that a' < a, b < b', and the interval [a', b'] also satisfies condition (f) of Theorem 1, then

$$\sup_{x \in [a,b]} \left| \int \frac{I_{[a',b']^c} d(F^{C_n}(\lambda) - F^{c_n,A_n,B_n}(\lambda))}{((x-\lambda)^2 + v_n^2)((x-\lambda)^2 + 2v_n^2) \cdots ((x-\lambda)^2 + 70v_n^2)} \right| + \sum_{\lambda_j \in [a',b']} \frac{v_n^{140}}{((x-\lambda_j)^2 + v_n^2)((x-\lambda_j)^2 + 2v_n^2) \cdots ((x-\lambda_j)^2 + 70v_n^2)} \right| = o(1), \text{ a.s.}$$
 (6)

where the  $\lambda_j$ 's denote the eigenvalues of  $C_n$ . From this, using the fact that the integral in (6) converges a.s. to 0, one can argue that, with probability one, no eigenvalue of  $C_n$  appears in [a, b] for all sufficiently large n.

Before proceeding, we simplify here some of the assumptions. It is clear from assumption (d) of Theorem 1 that, we can without loss of generality assume throughout that  $\max\{\|A_n\|, \|B_n\|\} \le 1$  for all n. Also, the argument given at the beginning of section 3 of Bai and Silverstein (1998) carries through in our case. Specifically, for any C > 0 let  $Y_{ij} = X_{ij}I_{(|X_{ij}| \le C)} - \mathbb{E}X_{ij}I_{(|X_{ij}| \le C)}$  (where  $I_A$  denotes the indicator function on the set A),  $Y_n = (Y_{ij})$ ,  $i \le n$ ,  $j \le N$ ,  $\tilde{C}_n = (1/N)A_n^{1/2}Y_nB_nY_n^*A_n^{1/2}$ , and  $\lambda_k$ ,  $\tilde{\lambda}_k$  the respective eigenvalues of  $C_n$  and  $\tilde{C}_n$  in nonincreasing order. Then as in Bai and Silverstein (1998), using the main result in Yin, Bai, and Krishnaiah (1988) on the largest eigenvalue of  $(1/N)X_nX_n^*$ , we have, with probability 1

$$\limsup_{n \to \infty} \max_{k \le n} |\lambda_k^{1/2} - \tilde{\lambda}_k^{1/2}| \le (1 + \sqrt{c}) \mathbb{E}^{1/2} |X_{11}|^2 I_{(|X_{11}| > C)},$$

and because of assumption (a) we can make the bound on the right side arbitrarily small by choosing C sufficiently large. Thus we can assume that the  $X_{ij}$  are uniformly bounded.

The rest of the paper is organized as follows. In Section 2, we give the key steps to the derivation of the integral equations for the limiting Stieltjes transforms of associated spectral measures. In Sections 3 and 4 we will show, respectively

$$\sup_{x \in [a,b]} |m_n(z) - \mathbb{E}m_n(z)| = o(1/(nv_n)) \quad \text{a.s.}$$
(7)

and

$$\sup_{x \in [a,b]} |\mathbb{E}m_n(z) - m_n^0(z)| = O(1/n). \tag{8}$$

Some mathematical tools needed in proving these results are given in the Appendix. Throughout this paper, K denotes a universal constant whose value may vary from one appearance to another.

# 2 Integral representation of Stieltjes transforms

Write  $X_n = [X_{\cdot 1}, \dots, X_{\cdot N}]$ , and let  $y_j = (1/\sqrt{N})A_n^{1/2}X_{\cdot j}$ . Then we can write

$$C_n = \sum_{j=1}^{N} b_j y_j y_j^*.$$

Fix  $z \in \mathbb{C}^+ \equiv \{z = x + iv \in \mathbb{C} : v > 0\}$ . Define

$$e_n = e_n(z) = (1/n)\operatorname{tr} A_n(C_n - zI)^{-1},$$
 (9)

and

$$p_n = -\frac{1}{Nz} \sum_{j=1}^{N} \frac{b_j}{1 + c_n b_j e_n} = \int \frac{-b}{z(1 + c_n b e_n)} dF^{B_n}(b).$$
 (10)

Write  $C_n = O\Lambda O^*$ ,  $\Lambda = \operatorname{diag}(\lambda_1, \dots, \lambda_n)$ , in its spectral decomposition. Let  $\underline{A}_n = \{\underline{a}_{ij}\} = O^* A_n O$ . Then

$$e_n = (1/n)\operatorname{tr}\underline{A}_n(\Lambda - zI)^{-1} = (1/n)\sum_{i=1}^n \frac{\underline{a}_{ii}}{\lambda_i - z}.$$
(11)

We therefore see that  $e_n$  is the Stieltjes transform of a measure on the nonnegative reals with total mass (1/n)tr  $A_n$ . It follows that both  $e_n(z)$  and  $ze_n(z)$  map  $\mathbb{C}^+$  into  $\mathbb{C}^+$ . This implies that  $p_n(z)$  and  $zp_n(z)$  map  $\mathbb{C}^+$  into  $\mathbb{C}^+$ , and as  $z \to \infty$ ,  $zp_n \to -(1/N)$ tr  $B_n$ . Therefore, from Lemma 5 we also have  $p_n$  a Stieltjes transform of a measure on the nonnegative reals with total mass (1/N)tr  $B_n$ . It follows that  $e_n$  and  $p_n$  are bounded in absolute value by  $v^{-1}(1/n)$ tr  $A_n$  and  $v^{-1}(1/N)$ tr  $B_n$ , respectively.

More generally, from Lemma 5 we have any function of the form

$$\frac{-b}{z(1+m(z))},$$

where  $b \ge 0$  and m(z) is the Stieltjes transform of a bounded measure on  $\mathbb{R}^+$ , to be the Stieltjes transform of a measure on the nonnegative reals with total mass b. It follows that

$$\left| \frac{-b}{z(1+m(z))} \right| \le \frac{b}{v}. \tag{12}$$

Let  $C_{(j)} = C_n - b_j y_j y_j^*$ . We may, without loss of generality, assume that  $\max(\|A_n\|, \|B_n\|) \le 1$ . Write

$$C_n - zI + zI + zp_n A_n = \sum_{j=1}^{N} b_j y_j y_j^* + zp_n A_n.$$

Taking inverses and using the definition of  $C_n$  and  $C_{(i)}$ , we have

$$(C_n - zI)^{-1} + (zI + zp_nA_n)^{-1}$$

$$= \sum_{j=1}^{N} b_j (C_n - zI)^{-1} y_j y_j^* (zI + zp_nA)^{-1} + zp_n (C_n - zI)^{-1} A_n (zI + zp_nA)^{-1}$$

$$= \sum_{j=1}^{N} b_j \frac{(C_{(j)} - zI)^{-1} y_j y_j^* (zI + zp_nA)^{-1}}{1 + b_j y_j^* (C_{(j)} - zI)^{-1} y_j} + zp_n (C_n - zI)^{-1} A_n (zI + zp_nA)^{-1},$$

where the last step follows from Lemma 1 in the appendix.

Taking traces and dividing by n, we have

$$m_n(z) - \int \frac{1}{a \int \frac{b}{1 + c_n b e_n} dF^{B_n}(b) - z} dF^{A_n}(a) = \frac{1}{N} \sum_{j=1}^N b_j d_j \equiv w_n^m,$$

where

$$d_{j} = \frac{(1/n)x_{j}^{*}A_{n}^{1/2}(I + p_{n}A_{n})^{-1}(C_{(j)} - zI)^{-1}A_{n}^{1/2}x_{j}}{z(1 + b_{j}y_{i}^{*}(C_{(j)} - zI)^{-1}y_{j})} - \frac{(1/n)\operatorname{tr}(C_{n} - zI)^{-1}A_{n}(I + p_{n}A_{n})^{-1}}{z(1 + c_{n}b_{j}e_{n})}.$$

Multiplying both sides of the above matrix identity by  $A_n$ , and then taking traces and dividing by n, we find

$$e_n(z) - \int \frac{a}{a \int \frac{b}{1 + c_n b e_n} dF^{B_n}(b) - z} dF^{A_n}(a) = \frac{1}{N} \sum_{j=1}^N b_j d_j^e \equiv w_n^e,$$

where

$$d_j^e = \frac{(1/n)x_j^* A_n^{1/2} (I + p_n A_n)^{-1} A_n (C_{(j)} - zI)^{-1} A_n^{1/2} x_j}{z(1 + b_j y_j^* (C_{(j)} - zI)^{-1} y_j)} - \frac{(1/n) \operatorname{tr} A_n (C_n - zI)^{-1} A_n (I + p_n A_n)^{-1}}{z(1 + c_n b_j e_n)}.$$

#### 2.1 Bound on the approximation error

Notice that for each j,  $y_j^*(C_{(j)}-zI)^{-1}y_j$  can be viewed as a Stieltjes transform of a measure on  $\mathbb{R}^+$ . Therefore

$$\left| \frac{1}{z(1 + b_j y_j^* (C_{(j)} - zI)^{-1} y_j)} \right| \le \frac{1}{v}.$$

For each j, let  $e_{(j)} = e_{(j)}(z) = (1/n) \operatorname{tr} A_n (C_{(j)} - zI)^{-1}$ , and

$$p_{(j)} = p_{(j)}(z) = \int \frac{-b}{z(1 + c_n b e_{(j)})} dF^{B_n}(b),$$

both of course being Stieltjes transforms of measures on  $\mathbb{R}^+$ , along with the integrand for each b. Using Lemma 1 and Lemma 2(a) in the appendix, (12) and the fact that  $||B_n|| \le 1$ , we have

$$|p_n - p_{(j)}| = |e_n - e_{(j)}|c_n \left| \int \frac{b^2}{z(1 + c_n b e_n)(1 + c_n b e_{(j)})} dF^{B_n}(b) \right| \le \frac{4c_n^2}{nv^3}.$$
 (13)

In order to handle both  $w_n^m$ ,  $d_j$  and  $w_n^e$ ,  $d_j^e$  at the same time, we shall denote by  $E_n$  either  $A_n$  or  $I_n$ , and  $w_n$ ,  $d_j$  for now will denote either the original  $w_n^m$ ,  $d_j$  or  $w_n^e$ ,  $d_j^e$ .

Write  $d_j = d_j^1 + d_j^2 + d_j^3 + d_j^4$ , where

$$d_{j}^{1} = \frac{(1/n)x_{j}^{*}A_{n}^{1/2}(I+p_{n}A_{n})^{-1}E_{n}(C_{(j)}-zI)^{-1}A_{n}^{1/2}x_{j}}{z(1+b_{j}y_{j}^{*}(C_{(j)}-zI)^{-1}y_{j})} - \frac{(1/n)x_{j}^{*}A_{n}^{1/2}(I+p_{(j)}A_{n})^{-1}E_{n}(C_{(j)}-zI)^{-1}A_{n}^{1/2}x_{j}}{z(1+b_{j}y_{j}^{*}(C_{(j)}-zI)^{-1}y_{j})},$$

$$d_j^2 = \frac{(1/n)x_j^* A_n^{1/2} (I + p_{(j)} A_n)^{-1} E_n (C_{(j)} - zI)^{-1} A_n^{1/2} x_j}{z(1 + b_j y_j^* (C_{(j)} - zI)^{-1} y_j)} - \frac{(1/n) \operatorname{tr} E_n (C_{(j)} - zI)^{-1} A_n (I + p_{(j)} A_n)^{-1}}{z(1 + b_j y_j^* (C_{(j)} - zI)^{-1} y_j)},$$

$$d_j^3 = \frac{(1/n)\operatorname{tr} E_n(C_{(j)} - zI)^{-1}A_n(I + p_{(j)}A_n)^{-1}}{z(1 + b_jy_j^*(C_{(j)} - zI)^{-1}y_j)} - \frac{(1/n)\operatorname{tr} E_n(C_n - zI)^{-1}A_n(I + p_nA_n)^{-1}}{z(1 + b_jy_j^*(C_{(j)} - zI)^{-1}y_j)},$$

and

$$d_j^4 = \frac{(1/n)\operatorname{tr} E_n(C_n - zI)^{-1}A_n(I + p_nA_n)^{-1}}{z(1 + b_jy_j^*(C_{(j)} - zI)^{-1}y_j)} - \frac{(1/n)\operatorname{tr} E_n(C_n - zI)^{-1}A_n(I + p_nA_n)^{-1}}{z(1 + c_nb_je_n)}.$$

For the rest of this subsection, we assume that for all n large,  $v = v_n = \kappa n^{-\delta}$  for some  $\kappa > 0$  and  $\delta \ge 0$ . We wish to show that for all  $\delta \le 1/4$ , for arbitrary subset  $S_n \subset \mathbb{R}$  containing at most n elements, and arbitrary positive t and  $\epsilon$ , we have

$$\mathbb{P}(\max_{x \in S_n} |w_n| v_n^{-12} > \epsilon) \le K_t \epsilon^{-2t} n^{2-t(1-34\delta)}$$

$$\tag{14}$$

by proving the same bound on each of

$$\mathbb{P}(\max_{j \le N, x \in S_n} |d_j^i| v_n^{-12} > \epsilon),$$

for i = 1, 2, 3, 4. Note that the constants  $K_t$  and  $K_p$  (appearing later) are positive constants. We begin with  $d_j^1$ . We get from Lemma 2(c) and (13),

$$|d_j^1| \leq \frac{1}{v_n} \frac{4c_n^2}{nv_n^3} \frac{1}{v_n} \frac{\|X_{\cdot j}\|^2}{n} \frac{16}{v_n^2} = \frac{64c_n^2}{nv_n^7} \frac{\|X_{\cdot j}\|^2}{n}.$$

From Lemma 3 it is straightforward to argue that for  $p \geq 2$ 

$$\mathbb{E}||X_{\cdot j}||^{2p} \le K_p n^p.$$

Therefore for  $p \geq 2$ 

$$\mathbb{P}(\max_{j \le N, x \in S_n} |d_j^1| v_n^{-12} > \epsilon) \le n \mathbb{P}\left(\max_{j \le N} \frac{\|X_{\cdot j}\|^2 64c_n^2}{n^2 v_n^{19}} > \epsilon\right) \le K_p \frac{nN}{(nv_n^{19})^p} \epsilon^{-p}.$$

For  $d_i^2$  we use Lemma 2(a) and Lemma 3 to get, for  $p \ge 2$ ,

$$\mathbb{E}|v_n^{-12}d_j^2|^p \le K_p v_n^{-13p} n^{-p/2} v_n^{-2p} = K_p \frac{1}{(n^{1/2} v_n^{15})^p},$$

so that for  $\epsilon > 0$ ,  $p \geq 2$ 

$$\mathbb{P}(\max_{j \le N, x \in S_n} |d_j^2| v_n^{-12} > \epsilon) \le K_p \epsilon^{-p} \frac{nN}{(n^{1/2} v_n^{15})^p}.$$

Using Lemma 1, Lemma 2(a), 2(b), and (13) we have,

$$|d_j^3 v_n^{-12}| \le \frac{K}{v_n^{13}} \left( \frac{1}{nv_n^2} + \frac{1}{nv_n^6} \right) \le K \frac{1}{nv_n^{19}}.$$

Therefore for any  $p \ge 1$  and  $\epsilon > 0$ 

$$\mathbb{P}(\max_{j \le N, x \in S_n} |d_j^3| v_n^{-12} > \epsilon) \le K_p \epsilon^{-p} \frac{nN}{(nv_n^{19})^p}.$$

Finally, for  $d_i^4$ , we use Lemma 1 and Lemma 2(a) to find:

$$|d_j^4 v_n^{-12}| \le \frac{4}{v_n^{16}} (|(1/n)X_{\cdot j}^* A_n (C_{(j)} - zI)^{-1} X_{\cdot j} - (1/n) \operatorname{tr} A_n (C_{(j)} - zI)^{-1}| + (nv_n)^{-1}).$$

Therefore, by Lemma 3, for any  $\epsilon > 0$ ,  $p \geq 2$ , we have

$$\mathbb{P}(\max_{j \leq N, x \in S_n} |d_j^4| v_n^{-12} > \epsilon) \\
\leq \sum_{j \leq N, x \in S_n} \left[ \mathbb{P}\left(\frac{4}{v_n^{16}} |(1/n)X_{\cdot j}^* A_n (C_{(j)} - zI)^{-1} X_{\cdot j} - (1/n) \operatorname{tr} A_n (C_{(j)} - zI)^{-1}| > \frac{\epsilon}{2} \right) + K_p \epsilon^{-p} (n v_n^{11})^{-p} \right] \\
\leq K_p \epsilon^{-p} \frac{nN}{(n^{1/2} v_n^{17})^p},$$

which, for  $\delta \in [0, 1/4]$ , can easily be verified to be the largest of the four bounds. Therefore, (14) holds.

### 2.2 Existence, convergence, and continuity of the solution

We can at this stage provide a proof of the existence of a unique e with nonnegative imaginary part satisfying (2) for any z = x + iv, v > 0, and the a.s. convergence in distribution of  $F^{C_n}$  to F. We also show the continuous dependence of e on  $F^A$ ,  $F^B$ , and c. We see from (14) with  $\delta = 0$ ,  $\kappa = v$  and t > 3 we have

$$e_n(z) - \int \frac{a}{a \int \frac{b}{1 + c_n b e_n} dF^{B_n}(b) - z} dF^{A_n}(a)$$
 and  $m_n(z) - \int \frac{1}{a \int \frac{b}{1 + c_n b e_n} dF^{B_n}(b) - z} dF^{A_n}(a)$ 

converge a.s. to zero. Consider a realization for which both convergences to zero occur on a subsequence  $\{n_i\}$  for which  $e_n$  converges, say to e. Because of (12), we have, by the Dominated Convergence Theorem (DCT)

$$p_n = \int \frac{-b}{z(1+c_nbe_n)} dF^{B_n}(b) \to \int \frac{-b}{z(1+cbe)} dF^B(b) \equiv p$$

and

$$\int \frac{a}{a \int \frac{b}{1 + c_n b e_n} dF^{B_n}(b) - z} dF^{A_n}(a) = \int \frac{-a}{z(1 + c_n a p_n)} dF^{A_n}(a)$$

$$\rightarrow \int \frac{-a}{z(1 + c a p)} dF^{A}(a)$$

$$= \int \frac{a}{a \int \frac{b}{1 + c b e} dF^{B}(b) - z} dF^{A}(a)$$

along the subsequence. Thus e is a solution to (2).

From uniqueness of e, proved below, we must have convergence of  $e_n$  to e on the whole sequence. Therefore, again by the DCT we have

$$m_n(z) \to \int \frac{1}{a \int \frac{b}{1+cbe} dF^B(b) - z} dF^A(a).$$

This event occurs with probability one, and for a countable number of v's with a limit point. Since the limsup of the largest eigenvalue of  $C_n$  is a.s. bounded by  $(1+\sqrt{c})^2$  (by Lemma 4), the sequence  $\{F^{C_n}\}$  is almost surely tight. Therefore,  $F^{C_n}$  converges in distribution to F a.s.

For probability distribution functions  $F^{\underline{A}}$  and  $F^{\underline{B}}$  on [0,1] and  $\underline{c} > 0$ , let  $\underline{e} = \underline{e}(z)$  be a solution to (2) with  $F^A$ ,  $F^B$ , c replaced by  $F^{\underline{A}}$ ,  $F^{\underline{B}}$ , and  $\underline{c}$ , respectively. Assume that  $c \leq \underline{c}$ . Then we have

$$e - \underline{e} = \int \frac{a}{a \int \frac{b}{1 + \underline{c}b\underline{e}} dF^{\underline{B}}(b) - z} d(F^{A}(a) - F^{\underline{A}}(a))$$

$$+ \int \frac{a^{2} \int \frac{b}{1 + \underline{c}b\underline{e}} d(F^{\underline{B}}(b) - F^{B}(b))}{(a \int \frac{b}{1 + \underline{c}b\underline{e}} dF^{B}(b) - z)(a \int \frac{b}{1 + \underline{c}b\underline{e}} dF^{\underline{B}}(b) - z)} dF^{A}(a)$$

$$+ (c - \underline{c}) \int \underline{e} \frac{a^{2} \int \frac{b^{2}}{(1 + \underline{c}b\underline{e})(1 + \underline{c}b\underline{e})} dF^{B}(b)}{(a \int \frac{b}{1 + \underline{c}b\underline{e}} dF^{B}(b) - z)(a \int \frac{b}{1 + \underline{c}b\underline{e}} dF^{\underline{B}}(b) - z)} dF^{A}(a) + \gamma(e - \underline{e}), \quad (15)$$

where

$$\gamma = c \int \frac{a^2 \int \frac{b^2}{(1+cbe)(1+\underline{c}b\underline{e})} dF^B(b)}{(a \int \frac{b}{1+cbe} dF^B(b) - z)(a \int \frac{b}{1+cbe} dF^B(b) - z)} dF^A(a).$$

Notice that by (12), the first integrand in (15) is bounded in absolute value by 1/v, the second by  $|z|/v^3$ , and the third by  $|z|^2/v^5$ . Let  $e_2$  and  $\underline{e}_2$  denote the imaginary parts of e and  $\underline{e}$ . Then we write

$$e_2 = \int \frac{a(ace_2 \int \frac{b^2}{|1 + cbe|^2} dF^B(b) + v)}{|a \int \frac{b}{1 + cbe} dF^B(b) - z|^2} dF^A(a) = e_2 \alpha + v\beta,$$

and

$$\underline{e}_2 = \int \frac{a(a\underline{c} \ \underline{e}_2 \int \frac{b^2}{|1+\underline{c}b\underline{e}|^2} dF^{\underline{B}}(b) + v)}{|a \int \frac{b}{1+\underline{c}b\underline{e}} dF^{\underline{B}}(b) - z|^2} dF^{\underline{A}}(a) = \underline{e}_2 \underline{\alpha} + v \underline{\beta}.$$

By Cauchy-Schwarz inequality:

$$\begin{split} |\gamma| & \leq \int \left( \frac{ca^2 \int \frac{b^2}{|1+cbe|^2} dF^B(b)}{\left| a \int \frac{b}{1+cbe} dF^B(b) - z \right|^2} \right)^{1/2} \left( \frac{\underline{c}a^2 \int \frac{b^2}{|1+\underline{c}b\underline{e}|^2} dF^B(b)}{\left| a \int \frac{b}{1+\underline{c}b\underline{e}} dF^B(b) - z \right|^2} \right)^{1/2} dF^A \\ & \leq \left( \int \frac{ca^2 \int \frac{b^2}{|1+cbe|^2} dF^B(b)}{\left| a \int \frac{b}{1+cbe} dF^B(b) - z \right|^2} dF^A(a) \right)^{1/2} \left( \int \frac{\underline{c}a^2 \int b^2 |1 + \underline{c}b\underline{e}|^2 dF^B(b)}{\left| a \int \frac{b}{1+\underline{c}b\underline{e}} dF^B(b) - z \right|^2} dF^A(a) \right)^{1/2} \\ & = \left( \frac{e_2 \alpha}{e_2 \alpha + v \beta} \right)^{1/2} \left( \frac{\underline{e}_2 \underline{\alpha}}{\underline{e}_2 \underline{\alpha} + v \underline{\beta}} \right)^{1/2}. \end{split}$$

Notice that for v small we have by Lemma 2(a)

$$e_2\alpha/\beta \le e_2c \int \frac{b^2}{|1+cbe|^2} dF^B(b) = -\Im \int \frac{b}{1+cbe} dF^B \le \frac{4c}{v}.$$

Therefore

$$\left( \int \frac{ca^2 \int \frac{b^2}{|1 + cbe|^2} dF^B(b)}{\left| a \int \frac{b}{1 + cbe} dF^B(b) - z \right|^2} dF^A(a) \right)^{1/2} = \left( \frac{e_2 \alpha}{e_2 \alpha + v \beta} \right)^{1/2}$$

$$= \left(\frac{e_2 \alpha / \beta}{v + e_2 \alpha / \beta}\right)^{1/2} \le \left(\frac{4c}{v^2 + 4c}\right)^{1/2} \le 1 - Kv^2$$

for v small, and for some positive constant K. A corresponding bound obviously exists for the other factor making up the bound on  $\gamma$ , so we conclude that for v small

$$|\gamma| \le 1 - Kv^2 \tag{16}$$

for some positive constant K.

We see then that (15) and (16) together reveal two things: uniqueness of solutions to (2) (with  $F^A = F^{\underline{A}}$ ,  $F^B = F^{\underline{B}}$ , and  $c = \underline{c}$ ), and continuous dependence of solutions to (2) on  $F^A$ ,  $F^B$  under the topology of weak convergence of probability measures (which follows from the DCT), and c.

#### 2.3 Bound on the difference between Stieltjes transforms

At this point on we assume that  $v_n = \kappa n^{-\delta}$  with  $\delta \in (0, 1/35]$ , so that  $v_n \to 0$  as  $n \to \infty$ . The main goal in this subsection is to prove the following:

$$\mathbb{P}(\max_{x \in S_n} v_n^{-1} | m_n - m_n^0 | > \epsilon) \le K_t \epsilon^{-t} n^{-\delta t/4} \quad \text{for } t \ge 280.$$
 (17)

We have  $e_n^0 = e_n^0(z)$  (with z = x + iv where v > 0 is arbitrary) a unique solution to

$$e = \int \frac{a}{a \int \frac{b}{1 + c_n b e} dF^{B_n}(b) - z} dF^{A_n}(a).$$
 (18)

Let  $m_n^0$  denote the Stieltjes transform of  $F^{c_n,A_n,B_n}$ . Then,

$$m_n^0 = m_n^0(z) = \int \frac{1}{a \int \frac{b}{1 + c_n b e_n^0} dF^{B_n}(b) - z} dF^{A_n}(a).$$
 (19)

Let  $e_2^0$ ,  $e_2$ ,  $m_2^0$ ,  $m_2$  denote the imaginary parts of  $e_n^0$ ,  $e_n$ ,  $m_n^0$ ,  $m_n$ , respectively. Also, let  $w_n^e$  and  $w_n^m$  be as defined earlier in Section 2.1. Then

$$e_{2}^{0} = \int \frac{a(ac_{n}e_{2}^{0}\int \frac{b^{2}}{|1+c_{n}be_{n}^{0}|^{2}}dF^{B_{n}}(b) + v)}{|a\int \frac{b}{1+c_{n}be_{n}^{0}}dF^{B_{n}}(b) - z|^{2}}dF^{A_{n}}(a),$$

$$e_{2} = \int \frac{a(ac_{n}e_{2}\int \frac{b^{2}}{|1+c_{n}be_{n}|^{2}}dF^{B_{n}}(b) + v)}{|a\int \frac{b}{1+c_{n}be_{n}}dF^{B_{n}}(b) - z|^{2}}dF^{A_{n}}(a) + \Im w_{n}^{e},$$

$$m_{2}^{0} = \int \frac{ac_{n}e_{2}^{0}\int \frac{b^{2}}{|1+c_{n}be_{n}^{0}|^{2}}dF^{B_{n}}(b) + v}{|a\int \frac{b}{1+c_{n}be_{n}^{0}}dF^{B_{n}}(b) - z|^{2}}dF^{A_{n}}(a),$$

$$m_{2} = \int \frac{ac_{n}e_{2}\int \frac{b^{2}}{|1+c_{n}be_{n}|^{2}}dF^{B_{n}}(b) + v}{|a\int \frac{b}{1+c_{n}be_{n}}dF^{B_{n}}(b) - z|^{2}}dF^{A_{n}}(a) + \Im w_{n}^{m},$$

and as above we have

$$e_n - e_n^0 = (e_n - e_n^0)\gamma_n + w_n^e,$$

where

$$|\underline{\gamma}_n| \leq \left( \int \frac{c_n a^2 \int \frac{b^2}{|1 + c_n b e_n|^2} dF^{B_n}(b)}{\left| a \int \frac{b}{1 + c_n b e_n} dF^{B_n}(b) - z \right|^2} dF^{A_n}(a) \right)^{1/2} \left( \int \frac{c_n a^2 \int \frac{b^2}{|1 + c_n b e_n^0|^2} dF^{B_n}(b)}{\left| a \int \frac{b}{1 + c_n b e_n^0} dF^{B_n}(b) - z \right|^2} dF^{A_n}(a) \right)^{1/2}.$$

Using the argument leading up to (16), we have for small v > 0,

$$\left(\int \frac{c_n a^2 \int \frac{b^2}{|1 + c_n b e_n^0|^2} dF^{B_n}(b)}{\left| a \int \frac{b}{1 + c_n b e_n^0} dF^{B_n}(b) - z \right|^2} dF^{A_n}(a) \right)^{1/2} \le \left(\frac{4c_n}{v^2 + 4c_n}\right)^{1/2} \le 1 - Kv^2$$

for some positive constant K.

Let  $\mu_n = n^{\delta/4}$ . Therefore we have  $v_n \mu_n^3 \to 0$ .

Let  $\lambda_{\max}$  denote the largest eigenvalue of  $(1/N)X_nX_n^*$ , and let  $K_1>(1+\sqrt{c})^2$ . Since

$$|\lambda - (x + iv_n)|^{-1} \le \frac{3}{2}|x|^{-1}$$

for fixed  $\lambda$  and |x| large, we have for all n large

$$|e_n - e_n^0| \le 3\mu_n^{-1} v_n^3 \tag{20}$$

whenever  $|x| > \mu_n v_n^{-3}$  and  $\lambda_{\max} \leq K_1$ . Notice that, since  $\max\{|e_n|, |e_n^0|\} \leq v_n^{-1} \frac{1}{n} \operatorname{tr} A_n$ , whenever  $(1/n)\operatorname{tr} A_n \leq v_n^4 \mu_n^{-1}$  we have

$$v_n^{-3}|e_n - e_n^0| \le 2v_n^{-4}(1/n)\operatorname{tr} A_n \le 2\mu_n^{-1}.$$
 (21)

Now, let  $\alpha$ ,  $\beta$  be such that  $e_2 = e_2\alpha + v_n\beta + \Im w_n^e$ . Then

$$\int \frac{c_n a^2 \int \frac{b^2}{|1 + c_n b e_n|^2} dF^{B_n}(b)}{\left| a \int \frac{b}{1 + c_n b e_n} dF^{B_n}(b) - z \right|^2} dF^{A_n}(a) = \frac{e_2 \alpha}{e_2 \alpha + v_n \beta + \Im w_n^e}.$$

Using Cauchy-Schwarz we have

$$|e_n| \le \beta^{1/2} ((1/n) \operatorname{tr} A_n)^{1/2} + |w_n^e|.$$

Next, as in Bai and Silverstein (1998) p. 329, for all n large  $|e_n| \ge \frac{1}{2}\mu_n^{-1}v_n^3(1/n)\mathrm{tr}\,A_n$  whenever  $|x| \le \mu_n v_n^{-3}$  and  $\lambda_{\max} \le K_1$ . So, for all n large, whenever  $|x| \le \mu_n v_n^{-3}$ ,  $|w_n^e| \le v_n^{12}$ ,  $\lambda_{\max} \le K_1$ , and  $(1/n)\mathrm{tr}\,A_n > v_n^4\mu_n^{-1}$  we have

$$\frac{1}{2}\mu_n^{-1}v_n^3(1/n)\operatorname{tr} A_n \le |e_n| \le \beta^{1/2}((1/n)\operatorname{tr} A_n)^{1/2} + |w_n^e| \le \beta^{1/2}((1/n)\operatorname{tr} A_n)^{1/2} + \mu_n v_n^8(1/n)\operatorname{tr} A_n.$$

Therefore

$$\frac{1}{3}\mu_n^{-1}v_n^3(1/n)\operatorname{tr} A_n \le (1/n)\operatorname{tr} A_n(\frac{1}{2}\mu_n^{-1}v_n^3 - \mu_n v_n^8) \le \beta^{1/2}((1/n)\operatorname{tr} A_n)^{1/2},$$

from which we get

$$\beta \ge \frac{1}{9} v_n^{10} \mu_n^{-3}.$$

Therefore

$$v_n\beta + \Im w_n^e \ge \frac{1}{9}v_n^{11}\mu_n^{-3} - v_n^{12} > 0,$$

and so

$$|e_n - e_n^0| \le K^{-1} v_n^{-2} |w_n^e| \le K^{-1} v_n^{10}$$
.

Thus, combining this with (21) and (20), we have, for all n large

$$\max_{x \in S_n} v_n^{-3} |e_n - e_n^0| \le K^{-1} v_n^7 + 3\mu_n^{-1} + 2v_n^{-4} \max_{x \in S_n} (I_{[|w_n^e| > v_n^{12}]} + I_{[\lambda_{\max} > K_1]}).$$

Therefore, for these n, and for any positive  $\epsilon$  and t we have

$$\mathbb{P}(\max_{x \in S_n} v_n^{-3} | e_n - e_n^0 | > \epsilon)$$

$$\leq K_t \epsilon^{-t} \left( n^{-7\delta t} + n^{-\delta t/4} + v_n^{-4t} [\mathbb{P}(\max_{x \in S_n} | w_n^e | v_n^{-12} > 1) + \mathbb{P}(\lambda_{\max} > K_1)] \right)$$

$$\leq K_t \epsilon^{-t} n^{-\delta t/4}, \tag{22}$$

where the last step follows by replacing t with

$$\frac{\frac{17}{4}\delta t + 2}{1 - 34\delta}$$

in (14) and t with  $\frac{17}{4}\delta t$  in Lemma 4. Taking the difference between  $m_n$  and  $m_n^0$  and using Cauchy-Schwarz, we get

$$|m_{n} - m_{n}^{0}| \leq |e_{n} - e_{n}^{0}| \left( \int \frac{c_{n} \int \frac{b^{2}}{|1 + c_{n}be_{n}|^{2}} dF^{B_{n}}(b)}{\left|a \int \frac{b}{1 + c_{n}be_{n}} dF^{B_{n}}(b) - z\right|^{2}} dF^{A_{n}}(a) \right)^{1/2} \left( \int \frac{c_{n}a^{2} \int \frac{b^{2}}{|1 + c_{n}be_{n}^{0}|^{2}} dF^{B_{n}}(b)}{\left|a \int \frac{b}{1 + c_{n}be_{n}^{0}} dF^{B_{n}}(b) - z\right|^{2}} dF^{A_{n}}(a) \right)^{1/2} + |w_{n}^{m}|.$$

As before, we have the second factor on the right bounded above by 1, while by Lemma 2(a) the first factor is bounded above by  $4c_n^{3/2}v_n^{-2}$ . Therefore, for any positive  $\epsilon$  and t we get from (14) and (22),

$$\mathbb{P}(\max_{x \in S_n} v_n^{-1} | m_n - m_n^0 | > \epsilon) \leq \mathbb{P}(\max_{x \in S_n} v_n^{-3} | e_n - e_n^0 | > \epsilon (2c_n^{1/2})^{-1}) + \mathbb{P}(\max_{x \in S_n} | w_n^m | v_n^{-1} > \epsilon/2) \\
\leq K_t \epsilon^{-t} \max(n^{-\delta t/4}, n^{2-t(1/2-17\delta)}). \tag{23}$$

Now it is easy to verify (17) from (23).

## 2.4 A rate on $F^{C_n}$ outside the support

Let  $\mathbb{E}_0$  denote the expectation, and  $\mathbb{E}_k$  denote the conditional expectation with respect to the  $\sigma$ -field generated by  $\{y_1, \ldots, y_k\}$ . Also, let  $\underline{\epsilon} > 0$  be such that [a', b'], with  $a' = a - \underline{\epsilon}$  and  $b' = b + \underline{\epsilon}$ , also satisfy condition (f) of Theorem 1. The goal of this subsection is to prove the following bound:

$$\max_{k \le N} \mathbb{E}_k(F^{C_n}([a', b']))^2 = o(v_n^2), \text{ a.s.}$$
(24)

Suppose that the *n* elements in  $S_n$  are equally spaced between  $-\sqrt{n}$  and  $\sqrt{n}$ . Since, for  $|x_1 - x_2| \le 2n^{-1/2}$ ,

$$|m_n(x_1+iv_n)-m_n(x_2+iv_n)| \le 2n^{-1/2}v_n^{-2}$$
  
 $|m_n^0(x_1+iv_n)-m_n^0(x_2+iv_n)| \le 2n^{-1/2}v_n^{-2}$ 

and when  $|x| \ge \sqrt{n}$ , for n large enough, for  $K_1$  as in Lemma 4,

$$|m_n(x+iv_n)| \le 2n^{-1/2} + v_n^{-1}I_{[\lambda_{\max}>K_1]}$$

and

$$|m_n^0(x+iv_n)| \le 2n^{-1/2},$$

Therefore, for all n large

$$\sup_{x \in \mathbb{R}} |m_n - m_n^0| \le \max_{x \in S_n} |m_n - m_n^0| + 4n^{-1/2} v_n^{-2} + v_n^{-1} I_{[\lambda_{\max} > K_1]},$$

and hence we conclude from (17) and Lemma 4, that for these n, and for any  $\epsilon > 0$  and  $t \ge 280$ ,  $0 < \delta \le \frac{1}{35}$ ,

$$\mathbb{P}(v_n^{-1} \sup_{x \in \mathbb{R}} |m_n - m_n^0| > \epsilon) \le K_t \epsilon^{-t} (n^{-\delta t/4} + n^{-t(1/2 - 3\delta)}) \le K_t \epsilon^{-t} n^{-\delta t/4}. \tag{25}$$

Since for any r > 0,

$$\mathbb{E}_k(v_n^{-r}\sup_{x\in\mathbb{R}}|m_n(x+iv_n)-m_n^0(x+iv_n)|^r),$$

for k = 0, 1, ..., n forms a martingale, from Jensen's inequality it follows that for any  $t \ge 1$ ,

$$(\mathbb{E}_k(v_n^{-r}\sup_{x\in\mathbb{R}}|m_n(x+iv_n)-m_n^0(x+iv_n)|^r))^t$$

for k = 0, 1, ..., n forms a submartingale. Therefore, from Lemma 2.5 and Lemma 2.6 of Bai and Silverstein (1998), and (25), for any  $\epsilon > 0$ ,  $t \ge 1$ , and t > 0, so that  $2rt \ge 280$ , we have,

$$\mathbb{P}(\max_{k \leq N} \mathbb{E}_{k}(v_{n}^{-r} \sup_{x \in \mathbb{R}} |m_{n}(x + iv_{n}) - m_{n}^{0}(x + iv_{n})|^{r}) > \epsilon)$$

$$\leq \epsilon^{-t} \mathbb{E}(v_{n}^{-rt} \sup_{x \in \mathbb{R}} |m_{n}(x + iv_{n}) - m_{n}^{0}(x + iv_{n})|^{rt})$$

$$\leq 2\epsilon^{-t} K_{2rt}^{1/2} n^{-\delta rt/4}, \tag{26}$$

whenever  $\delta \in (0, 1/35]$ . From this, it follows that with probability 1,

$$\max_{k \le N} \mathbb{E}_k(v_n^{-r} \sup_{x \in \mathbb{R}} |m_n(x + iv_n) - m_n^0(x + iv_n)|^r) \to 0.$$
 (27)

Let  $\lambda_1 \geq \ldots \geq \lambda_n$  be the eigenvalues of  $C_n$ , and write

$$m_{nj} = m_{nj}^{out} + m_{nj}^{in}, \qquad j = 1, 2,$$

where j=1 refers to the real part of  $m_n$  and j=2 refers to the imaginary part of  $m_n$ , so that

$$m_{n2}^{out}(x+iv_n) = \frac{1}{N} \sum_{\lambda_j \in [a',b']} \frac{v_n}{(x-\lambda_j)^2 + v_n^2} ,$$

and

$$m_{n1}^{out}(x+iv) = \frac{1}{N} \sum_{\lambda_j \in [a',b']} \frac{x - \lambda_j}{(x - \lambda_j)^2 + v_n^2}.$$

Similarly, write  $m_{n1}^0$  and  $m_{n2}^0$  to mean the real and imaginary parts of  $m_n^0$ . By (27), with probability 1, we have

$$\max_{k \le N} \mathbb{E}_k(v_n^{-r} \sup_{x \in \mathbb{R}} |m_{n2}(x + iv_n) - m_{n2}^0(x + iv_n)|^r) \to 0.$$
 (28)

Define the sequence  $\{G_q\}_{q=1}^{\infty}$  of functions on  $\mathbb{R}^2$  by

$$G_{\sum_{i=1}^{n-1}(N(i)+1)+k}(x_1,x_2) = \mathbb{E}_k F^{C_n}(x_1) F^{C_n}(x_2),$$

for k = 0, 1, 2, ..., N(n). Clearly, each  $G_q$  is a probability distribution function on  $\mathbb{R}^2$ . Also, for  $q = \sum_{j=1}^{n-1} (N(j)+1)+k$ , the two-dimensional Stieltjes transform,  $m_q^{(G)}(x_1+iv_1,x_2+iv_2)$  of  $G_q$  is  $\mathbb{E}_k m_n(x_1+iv_1)m_n(x_2+iv_2)$ . Notice  $x < 0, \lambda > 0$  implies that

$$\left| \frac{1}{\lambda - (x + iv)} - \frac{1}{\lambda - x} \right| \le \frac{v}{x^2}.$$

Therefore, from (27), we have with probability 1,

$$|m_q^{(G)}(x_1, x_2) - m_n^0(x_1)m_n^0(x_2)| \to 0$$
, as  $q \to \infty$ ,

for countably many negative  $x_1$  having a negative limit point, and countably many negative  $x_2$  also having a negative limit point.

It is straightforward to show the following: Assume that  $f(z_1, z_2)$  is a function of two complex variables, and analytic on a open rectangle  $E_1 \times E_2 \subset \mathbb{C}^2$  (that is, for fixed  $z_1 \in E$   $f(z_1, z_2)$  is analytic in  $z_2$ , and visa versa). Let  $\{z_1^n\} \subset E_1$ ,  $\{z_2^n\} \subset E_2$ , where  $\{z_1^n\}$  has a limit point in  $E_1$ ,  $\{z_2^n\}$  has a limit point in  $E_2$ . Then f is uniquely determined by the values it places on the set  $\{(z_1, z_2) : z_1 \in \{z_1^n\}, z_2 \in \{z_2^n\}\}$ . This, together with the a.s. tightness of  $G_q$ , gives us, with probability 1,  $G_q(y_1, y_2)$  converging weakly to  $F(y_1)F(y_2)$ .

Notice that the integrands of

$$\int_{[a',b']^c \times [a',b']^c} \frac{d\mathbb{E}_k F^{C_n}(x_1) F^{C_n}(x_2)}{((x-x_1)^2 + v_n^2)((x-x_2)^2 + v_n^2)}$$

and

$$\int_{[a',b']^c} \frac{d\mathbb{E}_k F^{C_n}(x_1)}{(x-x_1)^2 + v_n^2}$$

on their respective domains are uniformly bounded and equicontinuous for  $x \in [a, b]$ . Therefore, from *Problem 8*, p. 17, in Billingsley (1968), and using the fact that

$$\int_{a'}^{b'} \frac{v_n}{(x-u)^2 + v_n^2} dF^{c_n, A_n, B_n}(u) = 0 \quad \text{for all } x \in \mathbb{R},$$

the sequence  $\{g_q\}_{q=1}^{\infty}$  defined by

$$g_q = \sup_{x \in [a,b]} \mathbb{E}_k v_n^{-2} |m_{n2}^{in}(x + iv_n) - m_{n2}^{0}(x + iv_n)|^2$$

for  $q = \sum_{j=1}^{n-1} (N(j)+1) + k$  (with  $0 \le k \le N(n)$ ), converges to 0 a.s. as  $n \to \infty$ . Thus we have, a.s.,  $\max_{n \ge n_0} \max_{1 \le k \le N(n)} g_{\sum_{j=1}^{n-1} (N(j)+1)+k} \to 0$ , as  $n_0 \to \infty$ . This implies that

$$\max_{0 \le k \le N} \sup_{x \in [a,b]} \mathbb{E}_k v_n^{-2} |m_{n2}^{in}(x+iv_n) - m_{n2}^0(x+iv_n)|^2 \to 0 \quad \text{a.s.}$$

This, together with (28), implies that

$$\sup_{x \in [a,b]} \max_{k \le N} v_n^{-2} \mathbb{E}_k(m_{n2}^{out}(x+iv_n))^2 \to 0.$$
 (29)

Now, for any  $x \in [a, b]$ , we have

$$v_n^{-1} m_{n2}^{out}(x+iv_n) \geq \int_a^b \frac{1}{(u-x)^2 + v_n^2} dF^{C_n}(u)$$

$$\geq \int_{[a,b] \cap [x-v_n, x+v_n]} \frac{1}{(u-x)^2 + v_n^2} dF^{C_n}(u)$$

$$\geq \frac{1}{2v_n^2} F^{C_n}([a,b] \cap [x-v_n, x+v_n]).$$

We select  $x_j \in [a, b]$ , j = 1, ..., J, such that  $v_n < x_j - x_{j-1}$ , and  $[a, b] \subset \bigcup_{j=1}^J [x_j - v_n, x_j + v_n]$ . Notice that, as a consequence,  $J \leq (b-a)v_n^{-1}$ . Then from the inequality above, it follows that, with probability 1,

$$v_{n}^{-2} \max_{k \leq N} \mathbb{E}_{k}(F^{C_{n}}([a,b]))^{2} \leq v_{n}^{-2} \max_{k \leq N} \mathbb{E}_{k}(\sum_{j=1}^{J} F^{C_{n}}([a,b] \cap [x_{j}-v_{n},x_{j}+v_{n}]))^{2}$$

$$\leq v_{n}^{-2} \max_{k \leq N} \mathbb{E}_{k}(\sum_{j=1}^{J} 2v_{n}m_{n2}^{out}(x_{j}+iv_{n}))^{2}$$

$$\leq 4J \max_{k \leq N} \sum_{j=1}^{J} \mathbb{E}_{k}(m_{n2}^{out}(x_{j}+iv_{n}))^{2}, \text{ by H\"older's inequality}$$

$$\leq 4J^{2} \max_{1 \leq j \leq J} \max_{k \leq N} \mathbb{E}_{k}(m_{n2}^{out}(x_{j}+iv_{n}))^{2}$$

$$\leq 4(b-a)^{2}v_{n}^{-2} \sup_{x \in [a,b]} \max_{k \leq N} \mathbb{E}_{k}(m_{n2}^{out}(x+iv_{n}))^{2}$$

$$\to 0.$$

by (29).

This shows that,

$$\max_{k \le N} \mathbb{E}_k(F^{C_n}([a, b]))^2 = o(v_n^2), \text{ a.s.}$$

Clearly, the same argument holds for [a', b'] replacing [a, b], and so we have (24). Now, taking  $\delta = 1/35$ , from (24) we get,

$$\max_{k \le N} \mathbb{E}_k(F^{C_n}([a', b']))^2 = o(N^{-2/35}) \text{ a.s.}$$
(30)

#### 3 Convergence of $m_n - \mathbb{E} m_n$

Throughout the rest of the paper we take  $v_n = \kappa n^{-\delta}$  with  $\delta = \frac{1}{140}$ , and some constant  $\kappa > 0$ . In this section, we verify (7). Since for real  $x_1, x_2, |m_n(x_1 + iv_n) - m_n(x_2 + iv_n)| \le |x_1 - x_2|v_n^{-2}$  (and from this, the same bound holds for  $|\mathbb{E}m_n(x_1+iv_n)-\mathbb{E}m_n(x_2+iv_n)|$ ), we can prove (7) if we prove

$$\max_{x \in S_n} Nv_n |m_n(x + iv_n) - \mathbb{E}m_n(x + iv_n)| \to 0, \quad \text{a.s. as } n \to \infty,$$
(31)

for the set  $S_n$  consisting of  $n^2$  points equally spaced in [a,b].

Write  $D = C_n - zI$ ,  $D_j = D - b_j y_j y_j^*$  (where  $y_j = \frac{1}{\sqrt{N}} A_n^{1/2} X_{\cdot j}$ ), and  $D_{jj'} = D_j - b_{j'} y_{j'} y_{j'}^*$ , for  $j' \neq j$ . Note that  $D_j = C_{(j)} - zI$ . Then,  $m_n = \frac{1}{n} tr D^{-1}$ . Also, let  $\overline{D} = C_n - \overline{z}I$ , where  $\overline{z}$  is the complex conjugate of z. Note that  $\overline{D} = D^*$ . Also, let

$$\alpha_{j} = y_{j}^{*} D_{j}^{-2} y_{j} - \frac{1}{N} \operatorname{tr} (D_{j}^{-2} A_{n}), \qquad a_{j} = \frac{1}{N} \operatorname{tr} (D_{j}^{-2} A_{n}),$$

$$\beta_{j} = \frac{1}{1 + b_{j} y_{j}^{*} D_{j}^{-1} y_{j}}, \qquad \hat{b}_{j} = \frac{1}{1 + b_{j} \frac{1}{N} \mathbb{E}[\operatorname{tr} (D_{j}^{-1} A_{n})]},$$

$$\gamma_{j} = y_{j}^{*} D_{j}^{-1} y_{j} - \frac{1}{N} \mathbb{E}[\operatorname{tr} (D_{j}^{-1} A_{n})], \qquad \hat{\gamma}_{j} = y_{j}^{*} D_{j}^{-1} y_{j} - \frac{1}{N} \operatorname{tr} (D_{j}^{-1} A_{n}).$$

We first derive bounds for the moments of  $\gamma_j$  and  $\hat{\gamma}_j$ . Integrating first with respect to  $X_{j}$ , that is, conditioning on the set  $\{X_{i}: j \neq i\}$ , and using Lemma 3, for all  $p \geq 2$ ,

$$\mathbb{E}|\hat{\gamma}_{j}|^{p} \leq K_{p} N^{-p} \mathbb{E}[\operatorname{tr}(A_{n}^{1/2} D_{j}^{-1} A_{n} \overline{D}_{j}^{-1} A_{n}^{1/2})]^{p/2} \leq K_{p} N^{-p/2} v_{n}^{-p}, \tag{32}$$

where the last step follows from the fact that  $||D_j^{-1}|| \le v_n^{-1}$ , and that  $||A_n|| \le 1$ . Now, using the fact that  $(\mathbb{E}_j - \mathbb{E}_{j-1})[f_n(X_{\cdot 1}, \dots, X_{\cdot N})]$  (for any bounded  $f_N$ ) forms a martingale difference sequence w.r.t. the sigma-fields  $\mathcal{F}_{j-1}$  generated by columns  $\{X_{\cdot 1}, \dots, X_{\cdot (j-1)}\}$ , and that  $\mathbb{E}_0[\operatorname{tr}(D_j^{-1}A_n)] = \mathbb{E}[\operatorname{tr}(D_j^{-1}A_n)], \text{ and } \mathbb{E}_N[\operatorname{tr}(D_j^{-1}A_n)] = \operatorname{tr}(D_j^{-1}A_n), \text{ from Burkholder's inequality}$ (Lemma 2.2 in Bai and Silverstein (1998))

$$\mathbb{E}|\gamma_{j} - \hat{\gamma}_{j}|^{p} = N^{-p} \mathbb{E} \left| \sum_{j'\neq j}^{N} (\mathbb{E}_{j'} - \mathbb{E}_{j'-1}) [\operatorname{tr}(D_{j}^{-1}A_{n})] \right|^{P} \\
= N^{-p} \mathbb{E} \left| \sum_{j'\neq j}^{N} \mathbb{E}_{j'} [\operatorname{tr}(D_{j}^{-1} - D_{jj'}^{-1})A_{n}] - \mathbb{E}_{j'-1} [\operatorname{tr}(D_{j}^{-1} - D_{jj'}^{-1})A_{n}] \right|^{p} \\
= N^{-p} \mathbb{E} \left| \sum_{j'\neq j}^{N} (\mathbb{E}_{j'} - \mathbb{E}_{j'-1}) \left[ \frac{b_{j}y_{j}^{*}D_{jj'}^{-1}A_{n}D_{jj'}^{-1}y_{j}}{1 + b_{j}y_{j}^{*}D_{jj'}^{-1}y_{j}} \right] \right|^{p} \\
\leq K_{p}N^{-p} \mathbb{E} \left( \sum_{j'\neq j} \left| (\mathbb{E}_{j'} - \mathbb{E}_{j'-1}) \left[ \frac{b_{j}y_{j}^{*}D_{jj'}^{-1}A_{n}D_{jj'}^{-1}y_{j}}{1 + b_{j}y_{j}^{*}D_{jj'}^{-1}y_{j}} \right] \right|^{2} \right)^{p/2} \\
\leq K_{p}N^{-p/2}v_{n}^{-p}, \tag{33}$$

where in the last step we use Lemma 2.10 of Bai and Silverstein (1998) to bound the term within conditional expectations by  $||A_n|| v_n^{-1} \leq v_n^{-1}$ .

Therefore, from (32) and (33) it follows that for any  $p \geq 2$ ,

$$\mathbb{E}|\gamma_j|^p \le K_p N^{-p/2} v_n^{-p}. \tag{34}$$

Next, we write

$$m_{n} - \mathbb{E}m_{n} = \frac{1}{n} \sum_{j=1}^{N} (\mathbb{E}_{j} - \mathbb{E}_{j-1}) [\operatorname{tr}(D^{-1})]$$

$$= -\frac{1}{n} \sum_{j=1}^{N} (\mathbb{E}_{j} - \mathbb{E}_{j-1}) \left[ b_{j} \frac{y_{j}^{*} D_{j}^{-2} y_{j}}{1 + b_{j} y_{j}^{*} D_{j}^{-1} y_{j}} \right] \quad (\operatorname{since} \mathbb{E}_{j} \operatorname{tr}(D_{j}^{-1}) = \mathbb{E}_{j-1} \operatorname{tr}(D_{j}^{-1}))$$

$$= -\frac{1}{n} \sum_{j=1}^{N} (\mathbb{E}_{j} - \mathbb{E}_{j-1}) \left[ b_{j} \frac{y_{j}^{*} D_{j}^{-2} y_{j}}{1 + b_{j} \frac{1}{N} \mathbb{E} [\operatorname{tr}(D_{j}^{-1} A_{n})]} \right]$$

$$+ \frac{1}{n} \sum_{j=1}^{N} (\mathbb{E}_{j} - \mathbb{E}_{j-1}) \left[ b_{j}^{2} \frac{y_{j}^{*} D_{j}^{-2} y_{j} (y_{j}^{*} D_{j}^{-1} y_{j} - \frac{1}{N} \mathbb{E} [\operatorname{tr}(D_{j}^{-1} A_{n})])^{2}}{(1 + b_{j} \frac{1}{N} \mathbb{E} [\operatorname{tr}(D_{j}^{-1} A_{n})])^{2}} \right]$$

$$- \frac{1}{n} \sum_{j=1}^{N} (\mathbb{E}_{j} - \mathbb{E}_{j-1}) \left[ b_{j}^{3} \frac{y_{j}^{*} D_{j}^{-2} y_{j} (y_{j}^{*} D_{j}^{-1} y_{j} - \frac{1}{N} \mathbb{E} [\operatorname{tr}(D_{j}^{-1} A_{n})])^{2}}{(1 + b_{j} y_{j}^{*} D_{j}^{-1} y_{j})} \right]$$

$$= - \frac{1}{n} \sum_{j=1}^{N} b_{j} \hat{b}_{j} \mathbb{E}_{j} \alpha_{j} + \frac{1}{n} \sum_{j=1}^{N} b_{j}^{2} \hat{b}_{j}^{2} \mathbb{E}_{j} a_{j} \hat{\gamma}_{j}$$

$$+ \frac{1}{n} \sum_{j=1}^{N} b_{j}^{2} \hat{b}_{j}^{2} (\mathbb{E}_{j} - \mathbb{E}_{j-1}) [\alpha_{j} \gamma_{j} - b_{j} y_{j}^{*} D_{j}^{-2} y_{j} \beta_{j} \gamma_{j}^{2}]$$

$$= W_{1} + W_{2} + W_{3}. \tag{35}$$

## 3.1 Boundedness of $\hat{b}_i$

Let

$$p_n^0 = -\frac{1}{z} \int \frac{b}{1 + c_n b e_n^0} dF^{B_n}(b)$$
 and  $\hat{p}_n = -\frac{1}{z} \int \frac{b}{1 + c_n b \mathbb{E}(e_n)} dF^{B_n}(b)$ .

We have

$$m_n^0 = -\frac{1}{z} \int \frac{1}{ap_n^0 + 1} dF^{A_n}(a), \quad \text{and} \quad e_n^0 = -\frac{1}{z} \int \frac{a}{ap_n^0 + 1} dF^{A_n}(a).$$
 (36)

We have then

$$e_n^0 = -\frac{1}{z} \frac{1}{p_n^0} \int \frac{ap_n^0 + 1 - 1}{ap_n^0 + 1} dF^{A_n}(a) = -\frac{1}{zp_n^0} - \frac{m_n^0}{p_n^0}.$$

Therefore

$$e_n^0 p_n^0 = -\frac{1}{z} - m_n^0.$$

Suppose  $z = z_j \in \mathbb{C}^+ \to x \in [a', b']$  as  $j \to \infty$ . Then

$$e_n^0(z)p_n^0(z) \to -\frac{1}{x} - m_n^0(x) \in \mathbb{R}.$$

We see that both  $\{p_n^0(z_j)\}$  and  $\{e_n^0(z_j)\}$  remain bounded, since if, say  $e_n^0$  goes unbounded on some subsequence,  $p_n^0$  would tend to zero on that subsequence, and from (36) it will render  $e_n^0$  converging

to a finite number, a contradiction. Since  $\Im m_n^0(x) = 0$ , we must have  $\lim_{j\to\infty} \Im p_n^0(z_j) = 0$ . This in turn implies  $\lim_{j\to\infty} \Im e_n^0(z_j) = 0$  as well. Therefore, the measures defining  $p_n^0$  and  $e_n^0$  have derivative 0 for each  $x \in [a',b']$ , so that (a',b') is outside the support of both these measures, which after considering a slightly larger  $\underline{\epsilon}$ , this statement extends to [a',b'].

From continuity, we have  $e_n^0(z) \to e^0(z)$ , and consequently,  $m_n^0 \to m^0(z)$ , and  $p_n^0(z) \to p^0(z)$ ,  $e^0$ ,  $m^0$ ,  $p^0$  defined for the limiting empirical distribution, for all  $z \in \mathbb{C}^+ \cup [a',b']$ . We must have  $[-1/p_n^0(a'), -1/p_n^0(b')]$  not intersecting with any of the eigenvalues of  $A_n$  (respectively,  $[-1/p^0(a'), -1/p^0(b')]$  not intersecting with the support of  $F^A$ ). Therefore, since  $p_n^0(x) \to p^0(x)$ , for x = a', a, b, b', and  $p^0(a') < p^0(a) < p^0(b) < p^0(b')$ , we must have  $-1/p_n^0(z)$  uniformly bounded away from the eigenvalues of  $A_n$  for all z = x + iv,  $x \in [a, b]$ , and for  $v \in [0, v_0]$  for some positive  $v_0$ .

Similarly,  $-1/(c_n e_n^0)$  is uniformly bounded away from the eigenvalues of  $B_n$  for all z = x + iv,  $x \in [a, b], v \in [0, v_0]$ . Therefore, using (22) and arguments analogous to those leading to (27) (now applied to  $e_n$  instead of  $m_n$ ), we have, with  $z = x + iv_n$ 

$$\sup_{x \in [a,b]} |\hat{p}_n(z) - p_n^0(z)| = \sup_{x \in [a,b]} |e_n^0(z) - \mathbb{E}(e_n(z))| \frac{c_n}{|z|} \left| \int \frac{b^2}{(1 + c_n b e_n^0)(1 + c_n b \mathbb{E}(e_n))} dF^{B_n}(b) \right|$$

$$\leq \sup_{x \in \mathbb{R}} \frac{K}{v_n} |e_n^0(z) - \mathbb{E}(e_n(z))| \leq K \mathbb{E}(v_n^{-1} \sup_{x \in \mathbb{R}} |e_n - e_n^0|) \to 0$$

as  $n \to \infty$ . Thus we conclude

$$\sup_{x \in [a,b]} \| (I + \hat{p}_n(z)A_n)^{-1} \| \le K, \tag{37}$$

and

$$\max_{j \le N} \sup_{x \in [a,b]} \frac{1}{|(1 + c_n b_j \mathbb{E}(e_n))|} \le K.$$
(38)

Let for  $j \neq j \leq N$ ,

$$\hat{b}_j = \frac{1}{1 + c_n b_j n^{-1} \mathbb{E} \operatorname{tr} (A_n D_j^{-1})} \quad \text{and} \quad \hat{b}_{j\underline{j}} = \frac{1}{1 + c_n b_j n^{-1} \mathbb{E} (\operatorname{tr} (A_n D_{jj}^{-1}))}.$$

From Lemma 1

$$|(1/n)\operatorname{tr}(A_nD_j^{-1}) - e_n| \le (nv_n)^{-1}$$
 and  $|(1/n)\operatorname{tr}(A_nD_{jj}^{-1}) - e_n| \le 2(nv_n)^{-1}$ ,

so that from (38) we also have for all n large

$$\max_{j \le N} \sup_{x \in [a,b]} (|\hat{b}_j|, \max_{\underline{j} \ne j} |\hat{b}_{j\underline{j}}|) \le K. \tag{39}$$

## **3.2** Bounds on $W_1, W_2, W_3$

Let  $F_{nj}$  be the spectral distribution of the matrix  $\sum_{k\neq j} b_k y_k y_k^*$ . From Lemma 2.12 of Bai and Silverstein (1998), and (30), we get

$$\max_{j} \mathbb{E}_{j}(F_{nj}([a',b']))^{2} = o(v_{n}^{8}) = o(N^{-2/35}), \text{ a.s.}$$
(40)

Define

$$\mathcal{B}_j = I_{[\mathbb{E}_{j-1}F_{nj}([a',b']) \leq v_n^4] \cap [\mathbb{E}_{j-1}(F_{nj}([a',b']))^2 \leq v_n^8]} \ .$$

Note that,  $\mathcal{B}_j = I_{[\mathbb{E}_j F_{nj}([a',b']) \leq v_n^4] \cap [\mathbb{E}_j (F_{nj}([a',b']))^2 \leq v_n^8]}$  a.s., and we have

$$\mathbb{P}\left(\cup_{j=1}^{N}[\mathcal{B}_{j}=0] \text{ i.o.}\right)=0.$$

Therefore, for any  $\varepsilon > 0$ ,

$$\mathbb{P}\left(\max_{x \in S_n} |Nv_n W_1| > \varepsilon \text{ i.o.}\right) \\
\leq \mathbb{P}\left(\left(\left[\max_{x \in S_n} \left| v_n \sum_{j=1}^N \mathbb{E}_j(\alpha_j) \right| > \underline{\varepsilon}\right] \bigcap_{j=1}^N [\mathcal{B}_j = 1]\right) \cup \left(\bigcup_{j=1}^N [\mathcal{B}_j = 0]\right) \text{ i.o.}\right) \\
\leq \mathbb{P}\left(\max_{x \in S_n} \left| v_n \sum_{j=1}^N \mathbb{E}_j(\alpha_j) \mathcal{B}_j \right| > \underline{\varepsilon} \text{ i.o.}\right),$$

where  $\underline{\varepsilon} = \inf_n n\varepsilon/(N \max_{1 \leq j' \leq n} b_{j'} |\hat{b}_{j'}|) > 0$ , since  $\max_{1 \leq j \leq N} |b_j| \leq 1$ , and  $\max_{1 \leq j \leq N} \sup_{x \in [a,b]} |\hat{b}_j|$  is bounded for all n (by (39)). Note that, for each  $x \in \mathbb{R}$ ,  $\{\mathbb{E}_j(\alpha_j)\mathcal{B}_j\}$  forms a martingale difference sequence.

By Lemma 2.1 in Bai and Silverstein (1998), and Lemma 3, for each  $x \in [a, b]$ , and  $p \ge 2$ ,

$$\mathbb{E}\left|v_{n}\sum_{j=1}^{N}\mathbb{E}_{j}(\alpha_{j})\mathcal{B}_{j}\right|^{p}$$

$$\leq K_{p}\left(\mathbb{E}\left(\sum_{j=1}^{N}\mathbb{E}_{j-1}|v_{n}\mathbb{E}_{j}(\alpha_{j})\mathcal{B}_{j}|^{2}\right)^{p/2} + \sum_{j=1}^{N}\mathbb{E}|v_{n}\mathbb{E}_{j}(\alpha_{j})\mathcal{B}_{j}|^{p}\right)$$

$$\leq K_{p}\left(\mathbb{E}\left(\sum_{j=1}^{N}\mathbb{E}_{j-1}v_{n}^{2}N^{-2}\mathcal{B}_{j}\operatorname{tr}\left(A_{n}^{1/2}D_{j}^{-2}A_{n}\overline{D}_{j}^{-2}A_{n}^{1/2}\right)\right)^{p/2} + v_{n}^{p}\sum_{j=1}^{N}\mathbb{E}|\alpha_{j}|^{p}\right)$$

$$\leq K_{p}v_{n}^{p}N^{-p}\mathbb{E}\left(\sum_{j=1}^{N}\mathcal{B}_{j}\mathbb{E}_{j-1}\operatorname{tr}\left(D_{j}^{-2}\overline{D}_{j}^{-2}\right)\right)^{p/2} \quad (\text{since } \|A_{n}\| \leq 1)$$

$$+ K_{p}v_{n}^{p}N^{-p}\sum_{j=1}^{N}\mathbb{E}\left(\operatorname{tr}\left(A_{n}^{1/2}D_{j}^{-2}A_{n}\overline{D}_{j}^{-2}A_{n}^{1/2}\right)\right)^{p/2} \quad (\text{by Lemma 3})$$

$$\leq K_{p}\left(v_{n}^{p}N^{-p}\mathbb{E}\left(\sum_{j=1}^{N}\mathcal{B}_{j}\mathbb{E}_{j-1}\operatorname{tr}\left(D_{j}^{-2}\overline{D}_{j}^{-2}\right)\right)^{p/2} + v_{n}^{-p}N^{1-p/2}\right),$$

since  $\max_{j} \parallel D_{j}^{-1} \parallel \leq v_{n}^{-1}$  and  $\parallel A_{n} \parallel \leq 1$ .

Let  $\lambda_{kj}$  denote the k-th smallest eigenvalue of  $\sum_{k'\neq j} b_{k'} y_{k'} y_{k'}^*$ . We have, for  $x \in [a,b]$ 

$$\sum_{j=1}^{N} \mathcal{B}_{j} \mathbb{E}_{j-1} \operatorname{tr} \left( D_{j}^{-2} \overline{D}_{j}^{-2} \right) \\
= \sum_{j=1}^{N} \mathcal{B}_{j} \mathbb{E}_{j-1} \left[ \sum_{\lambda_{kj} \notin [a',b']} \frac{1}{((\lambda_{kj} - x)^{2} + v_{n}^{2})^{2}} + \sum_{\lambda_{kj} \in [a',b']} \frac{1}{((\lambda_{kj} - x)^{2} + v_{n}^{2})^{2}} \right] \\
\leq \sum_{j=1}^{N} (n\underline{\epsilon}^{-4} + \mathcal{B}_{j} v_{n}^{-4} n \mathbb{E}_{j-1} F_{nj}([a',b'])) \\
\leq KN^{2}. \tag{41}$$

Here the last step follows from (40).

Therefore, for  $p \ge \frac{70}{34}$ ,

$$\mathbb{P}\left(\max_{x\in S_n}\left|v_n\sum_{j=1}^N\mathbb{E}_j(\alpha_j)\mathcal{B}_j\right|>\varepsilon\right)\leq K_{p,\varepsilon}n^2N^{-p/140},$$

which is summable when p > 420. Therefore, by Borel-Cantelli lemma,

$$\max_{x \in S_n} |W_1| = o(1/Nv_n) \text{ a.s.}$$
 (42)

Next we prove

$$\max_{x \in S_n} |W_2| = o(1/Nv_n) \text{ a.s.}$$
(43)

by following similar arguments. First, observing that  $\{\mathbb{E}_j(a_j\hat{\gamma}_j)\mathcal{B}_j\}$  forms a martingale difference sequence, and using Lemma 2.1 of Bai and Silverstein (1998), Lemmas 1 and 3, and the fact that  $|a_j| \leq \frac{n}{N}v_n^{-2}$ , we have,

$$\mathbb{E}\left|v_{n}\sum_{j=1}^{N}\mathbb{E}_{j}(a_{j}\hat{\gamma}_{j})\mathcal{B}_{j}\right|^{p}$$

$$\leq K_{p}\left(\mathbb{E}\left(\sum_{j=1}^{N}\mathbb{E}_{j-1}|v_{n}\mathbb{E}_{j}(a_{j}\hat{\gamma}_{j})\mathcal{B}_{j}|^{2}\right)^{p/2} + \sum_{j=1}^{N}\mathbb{E}|v_{n}\mathbb{E}_{j}(a_{j}\hat{\gamma}_{j})\mathcal{B}_{j}|^{p}\right)$$

$$\leq K_{p}v_{n}^{p}N^{-p}\mathbb{E}\left(\sum_{j=1}^{N}\mathcal{B}_{j}\mathbb{E}_{j-1}(|a_{j}|^{2}\mathrm{tr}\left(D_{j}^{-1}\overline{D}_{j}^{-1}\right))\right)^{p/2} \quad \text{(by Lemma 3)}$$

$$+ K_{p}v_{n}^{-p}N^{-p}\sum_{j=1}^{N}\left(\mathrm{tr}\left(D_{j}^{-1}\overline{D}_{j}^{-1}\right)\right)^{p/2} \quad \text{(by Lemma 3, and since } \max_{j}|a_{j}| \leq \frac{n}{N}v_{n}^{-2})$$

$$\leq K_{p}\left(v_{n}^{p}N^{-p}\mathbb{E}\left(\sum_{j=1}^{N}\mathcal{B}_{j}\mathbb{E}_{j-1}(|a_{j}|^{2}\mathrm{tr}\left(D_{j}^{-1}\overline{D}_{j}^{-1}\right)\right)\right)^{p/2} + v_{n}^{-2p}N^{1-p/2}\right),$$

since  $\max_j \parallel D_j^{-1} \parallel \leq v_n^{-1}$ , so that  $\max_j \operatorname{tr}(D_j^{-1}\overline{D}_j^{-1}) \leq nv_n^{-2}$ . Moreover, using same notation as before, the fact that  $\parallel A_n \parallel \leq 1$ , and arguing as in the derivation of (41), we have

$$\sum_{j=1}^{N} \mathcal{B}_{j} \mathbb{E}_{j-1}(|a_{j}|^{2} \operatorname{tr}(D_{j}\overline{D}_{j}^{-1}))$$

$$\leq \sum_{j=1}^{N} \mathcal{B}_{j} \mathbb{E}_{j-1} N^{-2} n \left[ \sum_{k} \frac{1}{((\lambda_{kj} - x)^{2} + v_{n}^{2})^{2}} \right] \left[ \sum_{k} \frac{1}{(\lambda_{kj} - x)^{2} + v_{n}^{2}} \right],$$
(by Cauchy-Schwarz applied to  $|a_{j}|^{2}$ )
$$\leq \sum_{j=1}^{N} \mathcal{B}_{j} N^{-2} n \mathbb{E}_{j-1} (n\underline{\epsilon}^{-4} + v_{n}^{-4} n F_{nj}([a', b'])) (n\underline{\epsilon}^{-2} + v_{n}^{-2} n F_{nj}([a', b']))$$

$$\leq KN^{2}.$$

Since  $\max_j \max\{b_j, \sup_{x \in [a,b]} |\hat{b}_j|\}$  is bounded, for large enough n, we have (43) by arguments similar to the ones used already in the derivation of (42).

Note that, Lemma 1 implies that

$$\max_{j} \sup_{x \in [a,b]} |b_j y_j^* D_j^{-2} y_j \beta_j| = \max_{j} \sup_{x \in [a,b]} \left| \frac{b_j y_j^* D_j^{-2} y_j}{1 + b_j y_j^* D_j^{-1} y_j} \right| \le \frac{1}{v_n}. \tag{44}$$

Using Lemma 2.2 of Bai and Silverstein (1998) followed by Hölder's inequality, we have

$$\mathbb{E} \left| v_n \sum_{j=1}^{N} (\mathbb{E}_j - \mathbb{E}_{j-1}) (\alpha_j \gamma_j - b_j y_j^* D_j^{-2} y_j \beta_j \gamma_j^2) \right|^p$$

$$\leq K_p v_n^p N^{p/2 - 1} \sum_{j=1}^{N} (\mathbb{E} |\alpha_j \gamma_j|^p + v_n^{-p} \mathbb{E} |\gamma_j|^{2p}) \quad \text{(by (44))}$$

$$\leq K_p v_n^p N^{p/2 - 1} \sum_{j=1}^{N} \left( N^{-p} \left( \mathbb{E} (\text{tr } (D_j^{-2} \overline{D}_j^{-2}))^p \right)^{1/2} N^{-p/2} v_n^{-p} + v_n^{-3p} N^{-p} \right)$$

$$\text{(by Cauchy-Schwarz, Lemma 3, the fact that } \| A_n \| \leq 1, \text{ and (34)}$$

$$\leq K_p v_n^p N^{p/2} (N^{-p} N^{p/2} v_n^{-2p} N^{-p/2} v_n^{-p} + v_n^{-3p} N^{-p}) \quad \text{(since } \| D_j^{-1} \| \leq v_n^{-1} )$$

$$\leq K_p v_n^{-2p} N^{-p/2}.$$

Thus, using arguments as in the proof of (42) and (43), we get

$$\max_{x \in S_n} |W_3| = o(1/Nv_n). \tag{45}$$

Hence, (31), and consequently, (7), follow from (42), (43) and (45).

#### 4 Convergence of expected value

In this section we prove (8). Recall the definitions of D,  $D_j$  and  $D_{jj}$  from Section 3. Also, let

$$\beta_{j\,\underline{j}} = \frac{1}{1 + b_{\underline{j}} y_{\underline{j}}^* D_{j\,\underline{j}}^{-1} y_{\underline{j}}}.$$

For  $j \neq \underline{j} \leq N$  let  $\lambda_{kj\underline{j}}$  denote the k-th smallest eigenvalue of  $D_{j\underline{j}}$ , and let  $F_{nj\underline{j}}$  denote the empirical distribution function of this matrix. Using (40) and Lemma 2.12 of Bai and Silverstein (1998) we get

$$\max_{j \neq j} \mathbb{E}(F_{nj\underline{j}}[a', b'])^2 = o(v_n^8).$$

Therefore

$$\max_{j \leq N} \sup_{x \in [a,b]} \mathbb{E}(\operatorname{tr} D_{j}^{-1} \overline{D}_{j}^{-1})^{2} = \max_{j \leq N} \sup_{x \in [a,b]} \mathbb{E}\left(\sum_{\lambda_{kj} \notin [a',b']} \frac{1}{(\lambda_{kj} - x)^{2} + v_{n}^{2}} + \sum_{\lambda_{kj} \in [a',b']} \frac{1}{(\lambda_{kj} - x)^{2} + v_{n}^{2}}\right)^{2}$$

$$\leq \max_{j \leq N} \sup_{x \in [a,b]} \mathbb{E}(n\underline{\epsilon}^{-2} + v_{n}^{-2} n F_{nj}([a',b']))^{2} \leq K n^{2},$$
and, 
$$\max_{j \leq N} \sup_{x \in [a,b]} \mathbb{E}(\operatorname{tr} D_{j}^{-2} \overline{D}_{j}^{-2})^{2} \leq \max_{j \leq N} \sup_{x \in [a,b]} \mathbb{E}(n\underline{\epsilon}^{-4} + v_{n}^{-4} n F_{nj}([a',b']))^{2} \leq K n^{2}.$$

The latter implies of course

$$\max_{j \le N} \sup_{x \in [a,b]} \mathbb{E} \mathrm{tr} \, D_j^{-2} \overline{D}_j^{-2} \le Kn.$$

Similarly,

$$\max_{j \neq \underline{j}} \sup_{x \in [a,b]} \mathbb{E}(\operatorname{tr} D_{j\underline{j}}^{-1} \overline{D}_{j\underline{j}}^{-1})^2 \leq Kn^2, \text{ and } \max_{j \neq \underline{j}} \sup_{x \in [a,b]} \mathbb{E}(\operatorname{tr} D_{j\underline{j}}^{-2} \overline{D}_{j\underline{j}}^{-2}) \leq Kn.$$

Moreover

$$\max_{j \neq \underline{j}} \sup_{x \in [a,b]} \mathbb{E}(\operatorname{tr} D_{j\underline{j}}^{-1} \overline{D}_{j\underline{j}}^{-1})^4 \leq \mathbb{E}(n\underline{\epsilon}^{-2} + v_n^{-2} n F_{nj\underline{j}}([a',b']))^4 \leq K n^4 (\underline{\epsilon}^{-8} + v_n^{-8} \mathbb{E}(F_{nj\underline{j}}([a',b']))^2) \leq K n^4.$$

Write

$$C_n - zI + zI + z\hat{p}_n A_n = \sum_{j=1}^{N} b_j y_j y_j^* + z\hat{p}_n A_n.$$

Taking first inverses and then expected values we have

$$\begin{split} & \mathbb{E}(C_n - zI)^{-1} + (zI + z\hat{p}_n A_n)^{-1} \\ &= \mathbb{E}\left[\sum_{j=1}^N b_j (C_n - zI)^{-1} y_j y_j^* (zI + z\hat{p}_n A_n)^{-1} + z\hat{p}_n D^{-1} A_n (zI + z\hat{p}_n A_n)^{-1}\right] \\ &= \sum_{j=1}^N b_j \left[\mathbb{E}\frac{(C_{(j)} - zI)^{-1} y_j y_j^* (zI + z\hat{p}_n A_n)^{-1}}{1 + b_j y_j^* D^{-1} y_j} - \frac{1}{z(1 + c_n b_j \mathbb{E}(e_n))} (\mathbb{E}(C_n - zI)^{-1} A_n (I + \hat{p}_n A_n)^{-1}\right]. \end{split}$$

Taking the trace on both sides and dividing by n we have

$$\mathbb{E}(m_n(z)) - \int \frac{1}{a \int \frac{b}{1 + c_n b \mathbb{E}(e_n)} dF^{B_n}(b) - z} dF^{A_n}(a) = \frac{1}{zN} \sum_{j=1}^N b_j \hat{d}_j \equiv \hat{w}_n^m,$$

where

$$\hat{d}_{j} = \mathbb{E}[\beta_{j}(1/n)X_{\cdot j}^{*}A_{n}^{1/2}(I+\hat{p}_{n}A_{n})^{-1}D_{j}^{-1}A_{n}^{1/2}X_{\cdot j}] - \frac{(1/n)\operatorname{tr}\mathbb{E}[D^{-1}]A_{n}(I+\hat{p}_{n}A_{n})^{-1}}{(1+c_{n}b_{j}\mathbb{E}(e_{n}))}.$$

Multiplying both sides of the above matrix identity by  $A_n$ , and then taking traces and dividing by n, we find

$$\mathbb{E}(e_n(z)) - \int \frac{a}{a \int \frac{b}{1 + c_n b \mathbb{E}(e_n)} dF^{B_n}(b) - z} dF^{A_n}(a) = \frac{1}{zN} \sum_{j=1}^N b_j \hat{d}_j^e \equiv \hat{w}_n^e,$$

where

$$\hat{d}_{j}^{e} = \mathbb{E}[\beta_{j}(1/n)X_{\cdot j}^{*}A_{n}^{1/2}(I+\hat{p}_{n}A_{n})^{-1}A_{n}D_{j}^{-1}A_{n}^{1/2}X_{\cdot j}] - \frac{(1/n)\operatorname{tr}A_{n}\mathbb{E}[D^{-1}]A_{n}(I+\hat{p}_{n}A_{n})^{-1}}{(1+c_{n}b_{j}\mathbb{E}(e_{n}))}.$$

Again, we let  $E_n$  denote either  $A_n$  or  $I_n$ . We first show that

$$n^{-1} \max_{j \le N} \sup_{x \in [a,b]} |(\operatorname{tr} E_n \mathbb{E}[D^{-1}] A_n (I + \hat{p}_n A_n)^{-1} - \operatorname{tr} E_n \mathbb{E}[D_j^{-1}] A_n (I + \hat{p}_n A_n)^{-1}| = O(n^{-1}).$$
 (46)

Using  $\beta_j = \hat{b}_j - b_j \beta_j \hat{b}_j \gamma_j$ , (3.3) of Bai and Silverstein (1998), (34), (37), and (39) we conclude that the left hand side of (46) becomes

$$= n^{-1} \max_{j \le N} b_j \sup_{x \in [a,b]} |\mathbb{E}[\beta_j y_j^* D_j^{-1} A_n (I + \hat{p}_n A_n)^{-1} E_n D_j^{-1} y_j]|$$

$$\le n^{-1} \max_{j \le N} b_j \sup_{x \in [a,b]} (|\hat{b}_j| |\mathbb{E}[y_j^* D_j^{-1} A_n (I + \hat{p}_n A_n)^{-1} E_n D_j^{-1} y_j]|$$

$$+ b_j |\hat{b}_j| |\mathbb{E}[\beta_j \gamma_j y^* D_j^{-1} A_n (I + \hat{p}_n A_n)^{-1} E_n D_j^{-1} y_j]| )$$

$$\le Kn^{-1} \max_{j \le N} \sup_{x \in [a,b]} (N^{-1} |\mathbb{E}[\operatorname{tr} A_n^{1/2} D_j^{-1} A_n (I + \hat{p}_n A_n)^{-1} E_n D_j^{-1} A_n^{1/2}]|$$

$$+ v_n^{-1} (\mathbb{E}|\gamma_j|^2)^{1/2} (\mathbb{E}|y_j^* D_j^{-1} A_n (I + \hat{p}_n A_n)^{-1} E_n D_j^{-1} y_j|^2)^{1/2}$$

$$\le Kn^{-1} \max_{j \le N} \sup_{x \in [a,b]} (N^{-1} \mathbb{E}[\operatorname{tr} D_j^{-1} \overline{D}_j^{-1}])$$

$$+ v_n^{-1} N^{-1/2} v_n^{-1} N^{-1} (\mathbb{E}[\operatorname{tr} D_j^{-2} \overline{D}_j^{-2}] + \mathbb{E}(\operatorname{tr} D_j^{-1} \overline{D}_j^{-1})^2)^{1/2} \le Kn^{-1}.$$

Thus (46) holds.

From Lemma 3 and (37) we get

$$\max_{j \leq N} \sup_{x \in [a,b]} \mathbb{E}|(1/n)x_{j}^{*}A_{n}^{1/2}(I+\hat{p}_{n}A_{n})^{-1}E_{n}D_{j}^{-1}A_{n}^{1/2}x_{j} - (1/n)\operatorname{tr} E_{n}D_{j}^{-1}A_{n}(I+\hat{p}_{n}A_{n})^{-1}|^{2}$$

$$\leq Kn^{-2} \max_{j \leq N} \sup_{x \in [a,b]} \mathbb{E}[\operatorname{tr} D_{j}^{-1}\overline{D}_{j}^{-1}] \leq Kn^{-1}.$$
(47)

We next show

$$\max_{j \le N} \sup_{x \in [a,b]} n^{-2} \mathbb{E} |\operatorname{tr} E_n D_j^{-1} A_n (I + \hat{p}_n A_n)^{-1} - \operatorname{tr} E_n \mathbb{E} [D_j^{-1}] A_n (I + \hat{p}_n A_n)^{-1}|^2 \le K n^{-1}.$$
 (48)

Using (3.3) of Bai and Silverstein (1998), and the fact that  $\beta_{j\underline{j}} = \hat{b}_{j\underline{j}} - b_{\underline{j}} \hat{b}_{j\underline{j}} \beta_{j\underline{j}} \gamma_{j\underline{j}}$ , the left hand

side of (48) becomes

$$\begin{split} &= \max_{j \leq N} \sup_{x \in [a,b]} n^{-2} \sum_{\underline{j} \neq j} \mathbb{E} | (\mathbb{E}_{\underline{j}} - \mathbb{E}_{\underline{j}-1}) \mathrm{tr} \, E_n D_j^{-1} A_n (I + \hat{p}_n A_n)^{-1} |^2 \\ &\leq 2 \max_{j \leq N} \sup_{x \in [a,b]} n^{-2} \sum_{\underline{j} \neq j} b_{\underline{j}}^2 \mathbb{E} | \beta_{j\underline{j}} y_{\underline{j}}^* D_{j\underline{j}}^{-1} A_n (I + \hat{p}_n A_n)^{-1} E_n D_{j\underline{j}}^{-1} y_{\underline{j}} |^2 \\ &= 2 \max_{j \leq N} \sup_{x \in [a,b]} n^{-2} \sum_{\underline{j} \neq j} b_{\underline{j}}^2 \mathbb{E} | (\hat{b}_{j\underline{j}} - b_{\underline{j}} \hat{b}_{j\underline{j}} \beta_{j\underline{j}} \gamma_{j\underline{j}}) y_{\underline{j}}^* D_{j\underline{j}}^{-1} A_n (I + \hat{p}_n A_n)^{-1} E_n D_{j\underline{j}}^{-1} y_{\underline{j}} |^2 \\ &\leq K \max_{j \leq N} \sup_{x \in [a,b]} n^{-2} \sum_{\underline{j} \neq j} \left[ \mathbb{E} | y_{\underline{j}}^* D_{j\underline{j}}^{-1} A_n (I + \hat{p}_n A_n)^{-1} E_n D_{j\underline{j}}^{-1} y_{\underline{j}} |^2 \right. \\ &+ v_n^{-2} (\mathbb{E} | \gamma_{j\underline{j}} |^4 \mathbb{E} | y_{\underline{j}}^* D_{j\underline{j}}^{-1} A_n (I + \hat{p}_n A_n)^{-1} E_n D_{j\underline{j}}^{-1} y_{\underline{j}} |^4)^{1/2} \right] \\ &\leq K \max_{j \leq N} \sup_{x \in [a,b]} n^{-2} \sum_{\underline{j} \neq j} b_{\underline{j}}^2 n^{-2} \left[ \mathbb{E} (\mathrm{tr} \, D_{j\underline{j}}^{-2} \overline{D}_{j\underline{j}}^{-2}) + \mathbb{E} (\mathrm{tr} \, D_{j\underline{j}}^{-1} \overline{D}_{j\underline{j}}^{-1})^4)^{1/2} \right] \\ &\leq K \max_{j \leq N} \sup_{x \in [a,b]} n^{-2} \sum_{\underline{j} \neq j} n^{-2} (\mathbb{E} (\mathrm{tr} \, D_{j\underline{j}}^{-2} \overline{D}_{j\underline{j}}^{-2})^2 + \mathbb{E} (\mathrm{tr} \, D_{j\underline{j}}^{-1} \overline{D}_{j\underline{j}}^{-1})^4)^{1/2} \right] \\ &\leq K \max_{j \leq N} \sup_{x \in [a,b]} n^{-2} \sum_{\underline{j} \neq j} n^{-2} (n + n^2 + v_n^{-4} n^{-1} (n^2 + n^4)^{1/2}) \\ &\leq K n^{-1}. \end{split}$$

So (48) is true.

We get the same bound when  $(I + \hat{p}_n A_n)^{-1}$  is removed from the expressions, that is, we also have

$$\max_{j \le N} \sup_{x \in [a,b]} \mathbb{E} |\gamma_j - \hat{\gamma}_j|^2 \le K n^{-1}.$$

Moreover, using (32),

$$\max_{j \le N} \sup_{x \in [a,b]} \mathbb{E} |\hat{\gamma}_j|^2 \le K n^{-2} \mathbb{E} [\operatorname{tr} D_j^{-1} \overline{D}_j^{-1}] \le K n^{-1}.$$

Thus

$$\max_{j \le N} \sup_{x \in [a,b]} \mathbb{E} |\gamma_j|^2 \le K n^{-1}.$$

Therefore, with  $\hat{d}_j^{em}$  denoting either  $\hat{d}_j$  or  $\hat{d}_j^e$ , and with  $\hat{w}^{em}$  denoting either  $\hat{w}^e$  or  $\hat{w}^m$ , we have,

using Lemma 1, (46), (47), (34), and the fact that  $\beta_j = \hat{b}_j - \hat{b}_j^2 \gamma_j + \hat{b}_j^2 \beta_j \gamma_j^2$ ,

$$\sup_{x \in [a,b]} |\hat{w}^{em}|$$

$$= \sup_{x \in [a,b]} \left| \frac{1}{zN} \sum_{j=1}^{N} b_j \hat{d}_j^{em} \right|$$

$$\leq Kn^{-1} + \max_{j \le N} \sup_{x \in [a,b]} \left| \mathbb{E}[\beta_j (1/n) x_j^* A_n^{1/2} (I + \hat{p}_n A_n)^{-1} E_n D_j^{-1} A_n^{1/2} x_j] - \frac{(1/n) \operatorname{tr} E_n \mathbb{E}[D_j^{-1}] A_n (I + \hat{p}_n A_n)^{-1}}{1 + c_n b_j \mathbb{E}(e_n)} \right|$$

$$\leq Kn^{-1} + K \max_{j \le N} \sup_{x \in [a,b]} \left( \mathbb{E}[\beta_j - \hat{b}_j + \frac{c_n b_j \hat{b}_j}{1 + c_n b_j \mathbb{E}(e_n)} \mathbb{E}(e_n - (1/n) \operatorname{tr} A_n^{1/2} D_j^{-1} A_n^{1/2}) |(1/n) (\mathbb{E}(\operatorname{tr} D_j^{-1} \overline{D}_j^{-1}))^{1/2} \right)$$

$$+ |\mathbb{E}[\beta_j ((1/n) x_j^* A_n^{1/2} (I + \hat{p}_n A_n)^{-1} E_n D_j^{-1} A_n^{1/2} x_j - (1/n) \operatorname{tr} E_n \mathbb{E}(D_j^{-1}) A_n (I + \hat{p}_n A_n)^{-1})]|)$$

$$\leq Kn^{-1} + K \max_{j \le N} \sup_{x \in [a,b]} \left( ||\hat{b}_j^2| \mathbb{E}[\gamma_j - \beta_j \gamma_j^2] + v_n^{-1} n^{-1} |n^{-1/2} + |\hat{b}_j|^2 |\mathbb{E}[(\gamma_j - \beta_j \gamma_j^2) ((1/n) x_j^* A_n^{1/2} (I + \hat{p}_n A_n)^{-1} E_n D_j^{-1} A_n^{1/2} x_j - (1/n) \operatorname{tr} E_n \mathbb{E}(D_j^{-1}) A_n (I + \hat{p}_n A_n)^{-1})]|\right)$$

$$\leq K(n^{-1} + \max_{j \le N} \sup_{x \in [a,b]} (\mathbb{E}[\gamma_j|^2 + v_n^{-2} \mathbb{E}[\gamma_j|^4)^{1/2}) n^{-1/2})$$

$$\leq K(n^{-1} + (n^{-1} + v_n^{-2} n^{-2} v_n^{-4})^{1/2} n^{-1/2} )$$

$$\leq K(n^{-1} + (n^{-1} + v_n^{-2} n^{-2} v_n^{-4})^{1/2} n^{-1/2} )$$

$$\leq Kn^{-1}.$$

$$(49)$$

As before, we have

$$\mathbb{E}(e_n) - e_n^0 = (\mathbb{E}(e_n) - e_n^0)\gamma_n + \hat{w}_n^e$$

where, after inserting  $\hat{p}_n$  and  $p_n^0$ 

$$|\gamma_n| \le \left( \int \frac{c_n a^2 \int \frac{b^2}{|1 + c_n b \mathbb{E}(e_n)|^2} dF^{B_n}(b)}{|z|^2 |a\hat{p}_n + 1|^2} dF^{A_n}(a) \right)^{1/2} \left( \int \frac{c_n a^2 \int \frac{b^2}{|1 + c_n b e_n^0|^2} dF^{B_n}(b)}{|z|^2 |ap_n^0 + 1|^2} dF^{A_n}(a) \right)^{1/2}.$$

$$(50)$$

Let  $G_n^0$ ,  $G^0$  denote the distribution functions defining  $e_n^0$ ,  $e^0$ . Then  $G_n^0 \xrightarrow{\mathcal{D}} G^0$ . We have

$$\int \frac{1}{(\lambda - x)^2} dG^0(\lambda) = \frac{d}{dx} e^0(x)$$

uniformly bounded for  $x \in [a, b]$ . For  $\lambda$  in either  $(-\infty, a']$  or  $[b', \infty)$ ,  $\{(\lambda - x)^{-2} : x \in [a, b]\}$  form a uniformly bounded, equicontinuous family of functions in  $\lambda$ . From Billingsley (1968), *Problem 8*, p. 17, we have then

$$\lim_{n \to \infty} \sup_{x \in [a,b]} \left| \frac{d}{dx} e_n^0(x) - \frac{d}{dx} e^0(x) \right| = 0.$$

Since for all  $x \in [a, b], \lambda \in [a', b']^c$  and positive v

$$\left| \frac{1}{(\lambda - x)^2 + v^2} + \frac{1}{(\lambda - x)^2} \right| \le \frac{v^2}{\underline{\epsilon}^4},$$

recalling that  $e_2^0 = \Im e_n^0$ , we have for any sequence of positive  $v_n'$  converging to 0,

$$\lim_{n \to \infty} \sup_{x \in [a,b]} \left| \frac{e_2^0(x + iv_n')}{v_n'} - \frac{d}{dx} e_n^0(x) \right| = 0.$$

Therefore, we conclude that

$$\sup_{n,x\in[a,b]} \frac{e_2^0(x+iv_n)}{v_n} \le K,\tag{51}$$

Writing again  $e_2^0 = e_2^0 \alpha + v_n \beta$  we have, by (51) and the conclusion in Section 3.1 concerning the eigenvalues of  $B_n$  remaining away from  $-1/(c_n e_n^0)$ ,

$$\sup_{x \in [a,b]} \frac{e_2^0 \alpha}{v_n \beta} \le \sup_{x \in [a,b]} \frac{e_2^0}{v_n} c_n \int \frac{b^2}{|1 + c_n b e_n^0|^2} dF^{B_n}(b) \le K.$$

Therefore

$$\sup_{x \in [a,b]} \int \frac{c_n a^2 \int \frac{b^2}{|1 + c_n b e_n^0|^2} dF^{B_n}(b)}{|z|^2 |ap_n^0 + 1|^2} dF^{A_n}(a) = \sup_{x \in [a,b]} \frac{e_2^0 \alpha / (v_n \beta)}{(e_2^0 \alpha / (v_n \beta)) + 1}$$

is uniformly bounded away from 1. Moreover, from continuity and the uniform convergence of  $\mathbb{E}(e_n)$  and  $\hat{p}_n$ , we must have that the supremum over all  $x \in [a, b]$  of the first factor on the right hand side of (50) is also uniformly bounded away from 1 for all n large. We therefore have from (49)

$$\sup_{x \in [a,b]} |\mathbb{E}(e_n) - e_n^0| = O(n^{-1}).$$

Again

$$|\mathbb{E}(m_{n}) - m_{n}^{0}| \leq |\mathbb{E}(e_{n}) - e_{n}^{0}| \left( \int \frac{c_{n} \int \frac{b^{2}}{|1 + c_{n}b\mathbb{E}(e_{n})|^{2}} dF^{B_{n}}(b)}{|z|^{2} |a\hat{p}_{n} + 1|^{2}} dF^{A_{n}}(a) \right)^{1/2} \left( \int \frac{c_{n}a^{2} \int \frac{b^{2}}{|1 + c_{n}be_{n}^{0}|^{2}} dF^{B_{n}}(b)}{|z|^{2} |ap_{n}^{0} + 1|^{2}} dF^{A_{n}}(a) \right)^{1/2} + |\hat{w}_{n}^{m}|.$$

The second factor on the right is of course bounded by 1, and from (37) and (38), the first factor is bounded, uniformly for  $x \in [a, b]$ . Therefore, by (49), we conclude that (8) holds.

Thus combining the results of this section and the previous section, we arrive at (3), and along with section 6 of Bai and Silverstein (1998), this completes the proof of Theorem 1.

# Appendix: mathematical tools

**Lemma 1** For  $n \times n$  A,  $\tau \in \mathbb{C}$ , and  $r \in \mathbb{C}^n$  for which A and  $A + \tau r r^*$  are invertible,

$$r^*(A + \tau r r^*)^{-1} = \frac{1}{1 + \tau r^* A^{-1} r} r^* A^{-1}.$$

(follows from  $r^*A^{-1}(A + \tau r r^*) = (1 + \tau r^*A^{-1}r)r^*$ .)

Moreover, (Lemma 2.6 of Silverstein and Bai (1995)), let  $z \in \mathbb{C}^+$  with v = Im z, A and B  $n \times n$  with B Hermitian, and  $r \in \mathbb{C}^n$ . Then

$$\left| \operatorname{tr} \left( (B - zI)^{-1} - (B + rr^* - zI)^{-1} \right) A \right| = \left| \frac{r^* (B - zI)^{-1} A (B - zI)^{-1} r}{1 + r^* (B - zI)^{-1} r} \right| \le \frac{\|A\|}{v}.$$

**Lemma 2** (Lemma 2.3 of Silverstein (1995)): For  $z = x + iv \in \mathbb{C}^+$  let  $m_1(z)$ ,  $m_2(z)$  be Stieltjes transforms of any two measures with respective total mass  $M_1$ ,  $M_2$ ; A, B, and C  $n \times n$  with A Hermitian nonnegative definite, and  $r \in \mathbb{C}^n$ . Then

(a) 
$$||(m_1(z)A + I)^{-1}|| \le \max(4M_1||A||/v, 2)$$

(b) 
$$|\operatorname{tr} B((m_1(z)A+I)^{-1} - (m_2(z)A+I)^{-1})| \le |m_2(z) - m_1(z)|n||B|| ||A|| \max(4M_1||A||/v, 2) \max(4M_2||A||/v, 2)$$

(c) 
$$|r^*B(m_1(z)A+I)^{-1}Cr - r^*B(m_2(z)A+I)^{-1}Cr$$

$$\leq |m_2(z) - m_1(z)|||r||^2 ||A|||B|| \max(4M_1||A||/v, 2) \max(4M_2||A||/v, 2)$$

(||r|| denoting Euclidean norm on r).

**Lemma 3** (Lemma 2.7 of Bai and Silverstein (1998)): For  $X = (X_{\cdot 1}, \dots, X_n)^T$  i.i.d. standardized and bounded entries,  $C \ n \times n$  matrix we have for any  $p \ge 2$ 

$$\mathbb{E}|X_{\cdot 1}^*CX_{\cdot 1} - \operatorname{tr} C|^p \le K_p(\operatorname{tr} CC^*)^{p/2}$$

where  $K_p$  depends on the distribution of  $X_{\cdot 1}$ .

**Lemma 4** (Analog of (3.1) of Bai and Silverstein (1998)): When the entries of  $X_n$  are bounded the largest eigenvalue of  $\frac{1}{N}X_nX_n^*$ , denoted by  $\lambda_{\max}$ , satisfies

$$\mathbb{P}(\lambda_{\max} > K) = o(n^{-t})$$

for any  $K > (1 + \sqrt{c})^2$  and any positive t.

**Lemma 5** (Lemma 2.2 of Shohat and Tamarkin (1970), and Theorems A.2, A.4, A.5 of Krein and Nudelman (1997)): If f is analytic on  $\mathbb{C}^+$ , both f(z) and zf(z) map  $\mathbb{C}^+$  into  $\mathbb{C}^+$ , and there is a  $\theta \in (0, \pi/2)$  for which  $zf(x) \to c$ , finite, as  $z \to \infty$  restricted to  $\{w \in \mathbb{C} : \theta < \arg w < \pi - \theta\}$ , then c < 0 and f is the Stieltjes transform of a measure on the nonnegative reals with total mass -c.

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